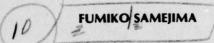
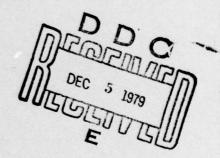


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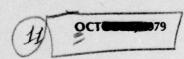
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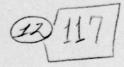




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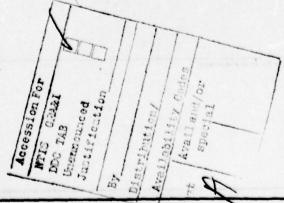
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Constant Information Model can be used when we substitute a subset of equivalent, binary items, whose item characteristic functions are unknown, for the "Old Test," or a subset of test items whose operating characteristics are known, in estimating the operating characteristics of item response. In so doing, it has been suggested that we choose items whose common discrimination power is low, so that the interval of latent trait, for which their common item characteristic function in the Constant Information Model assumes positive values, is wide enough to cover the ability levels of all the examinees, on which the operating characteristics of new items are to be estimated.

This suggestion needs more investigation and precision, however, since it is expected from theory that the convergence of the conditional distribution of the maximum likelihood estimate, given latent trait, to the normality is slow for the values of latent trait which are close to the two endpoints of the above interval, in comparison with those close to the midpoint. In this paper, through a simulation study, the speed of convergence of the conditional distribution of the maximum likelihood estimate to the normality at various levels of latent trait is observed, using twenty hypothetical test sessions, in each of which ten equivalent, binary test items are given. As was expected, the conditional distribution of the maximum likelihood estimate is skewed for the values of latent trait close to the two endpoints of the interval, and the convergence is slower. From these results, the above suggestion is added by more precision, which is essential to the researchers who wish to apply this method in estimating the operating characteristics.



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CONVERGENCE OF THE CONDITIONAL DISTRIBUTION OF THE MAXIMUM LIKELIHOOD ESTIMATE, GIVEN LATENT TRAIT, TO THE ASYMPTOTIC NORMALITY:

OBSERVATIONS MADE THROUGH THE CONSTANT INFORMATION MODEL

#### ABSTRACT

Constant Information Model can be used when we substitute a subset of equivalent, binary items, whose item characteristic functions are unknown, for the "Old Test," or a subset of test items whose operating characteristics are known, in estimating the operating characteristics of item response. In so doing, it has been suggested that we choose items whose common discrimination power is low, so that the interval of latent trait, for which their common item characteristic function in the Constant Information Model assumes positive values, is wide enough to cover the ability levels of all the examinees, on which the operating characteristics of new items are to be estimated.

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The research was conducted at the principal investigator's laboratory, 409 Austin Peay Hall, Department of Psychology, University of Tennessee, Knoxville, Tennessee. Those who worked in the laboratory and helped the author in various ways for this research include Paul S. Changas, Robert L. Trestman, Philip S. Livingston, and Lin Wen Kuei-Chiu.

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## I Introduction

Constant Information Model on the dichotomous response level (Samejima, 1972) of latent trait theory has been proposed and discussed (cf. Samejima, 1979). Let  $\theta$  be a unidimensional latent trait, whose range is given by

$$(1.1) \quad -\infty < \theta < \infty$$

and g be a binary item. The model is defined by the item characteristic function,  $P_{\bf g}(\theta)$  , such that

(1.2) 
$$P_g(\theta) = \sin^2[a_g(\theta-b_g) + (\pi/4)]$$
,

where  $a_g$  and  $b_g$  are the discrimination and difficulty parameters, respectively. The model provides us with such an item information function,  $I_g(\theta)$ , that

(1.3) 
$$I_{g}(\theta) \begin{cases} = 4a_{g}^{2} & \underline{\theta} < \theta < \overline{\theta} \\ = 0 & \text{otherwise} \end{cases},$$

where

(1.4) 
$$\frac{\theta}{2} = \left[-\pi a_g^{-1}/4\right] + b_g$$

and

(1.5) 
$$\bar{\theta} = [\pi a_g^{-1}/4] + b_g$$
.

When we consider a subinterval of  $\theta$  such that

$$(1.6) \qquad \underline{\theta} < \theta < \overline{\theta}$$

as its range, the item characteristic function of the Constant Information Model is strictly increasing in  $\theta$ , with zero and unity as its two asymptotes —— the characteristics shared by the normal ogive model, the logistic model, and so on. For convenience, hereafter, we shall call the set of models which share this particular characteristic Type A . It has been shown (Samejima, 1979) that the area under the curve of the square root of the item information function equals  $\pi$  regardless of the model, provided that it belongs to Type A .

One usefulness of the Constant Information Model is that, by virtue of the transformation-free character (Samejima, 1969) of the maximum likelihood estimate, we can adopt this model whenever we locate a suitable subset of equivalent, binary items in a newly developed item pool, and use this subset of items as a substitute for the set of items whose operating characteristics are known (cf. Samejima, 1979). Thus we can make use of various combinations of a method and an approach for estimating the operating characteristics of the item response categories (Samejima, 1977a, 1977b, 1978a, 1978b, 1978c, 1978d, 1978e, 1978f) without depending upon the subset of a priori investigated items, or the Old Test. The only assumption used in this procedure is that the "true" item characteristic function, which is common to these equivalent, binary items, belongs to Type A.

In the actual process of estimating the operating characteristics, a combination of a method and an approach is selected for use. In these combinations of a method and an approach, the asymptotic normality of the conditional distribution of the maximum likelihood estimate,  $\hat{\theta}$ , given  $\theta$ , plays an important role.

For the test of n equivalent, binary items, Constant

Information Model provides us with such a test information function that

(1.7) 
$$I(\theta) = 4na_g^2$$
,

for the interval of  $\theta$  given by (1.6). In the asymptotic normal distribution of the maximum likelihood estimate, the two parameters, i.e., the conditional mean and variance, are  $\theta$  and the inverse of the test information function.

Estimation of the operating characteristics of the graded item response categories will be pursued further, by making more use of the characteristics of the polynomials obtained by the method of moments (cf. Samejima and Livingston, 1979), and so forth. In so doing, the role of the Constant Information Model will be important, as long as we wish to use the methods for a newly developed item pool, without depending upon a set of "known" items, or the Old Test. One problem we should seriously investigate before using the Constant Information Model for a subset of "unknown," equivalent items of the item pool concerns the differences in the speed of convergence to the normality caused by the positions of  $\theta$  in the interval,  $(\underline{\theta}, \overline{\theta})$  (cf. Samejima, 1979). In the present paper, we shall pursue this subject through a Monte Carlo study.

### II Method and Data Calibration

For the common item characteristic function of the hypothetical equivalent, binary items, Constant Information Model with the parameters,

(2.1) 
$$\begin{cases} a_g = 0.25 \\ b_g = 0.00 \end{cases}$$

was used. The interval of  $\theta$  for which the item information function assumes a positive constant is, therefore, given from (1.4), (1.5) and (1.6) by

(2.2) 
$$-\pi < \theta < \pi$$
,

and we have for the amount of item information

(2.3) 
$$I_g(\theta) = 0.25$$
.

As the fixed levels of the latent trait  $\theta$ , eight positions were selected, i.e., -3.0, -2.2, -1.4, -0.6, 0.2, 1.0, 1.8 and 2.6. A group of one hundred hypothetical examinees were assigned to each of the eight levels of ability  $\theta$ , to make the total number of hypothetical examinees eight hundred. There were twenty hypothetical sessions of testing, and in each session ten equivalent, binary items were administered. An item score  $u_g$  (= 0 or 1) was calibrated by the Monte Carlo method following the Constant Information Model, whose parameters were given in the preceding paragraph. After the completion of each session, the (cumulative) test score t such that

(2.4) 
$$t = \sum_{\mathbf{u}_{\mathbf{g}} \in \mathbf{V}} \mathbf{u}_{\mathbf{g}} ,$$

where V is the response pattern obtained after the completion of each session, or the vector of item scores to the items, which were administered by the end of the session, was computed for each of the eight hundred hypothetical examinees. Thus after the completion of the k-th session the full test score is 10 x k . The maximum likelihood estimate  $\hat{\theta}$  was obtained by

(2.5) 
$$\hat{\theta} = P_g^{-1}[t/(10k)]$$
$$= 4 \sin^{-1}\{[t/(10k)]^{1/2}\} - \pi$$

for each hypothetical subject, after the completion of the k-th session (cf. Samejima, 1979).

Figure 2-1 illustrates three different transformations of  $\,\theta\,$  to  $\,\tau_1$  ,  $\,\tau_2$  and  $\,\tau_3$  , respectively, through the formulae such that

(2.6) 
$$\tau_1 = 5.00 \sin^2[0.25\theta + (\pi/4)] - 2.50$$
,

(2.7) 
$$\tau_2 = P_g^{\star\star^{-1}} \left[ \sin^2 \{0.25\theta + (\pi/4)\} \right] ,$$

where

I

-

(2.8) 
$$P_{g}^{**}(\theta) = (2\pi)^{-1/2} \int_{-\infty}^{s} \exp[-u^{2}/2] du ,$$

and

(2.9) 
$$\tau_3 = (2/1.7) \log [\tan\{(0.25\theta) + (\pi/4)\}]$$
.

They are the three transformations, which are eventually required if the true item characteristic function common to these equivalent items follows: 1) the linear model such that

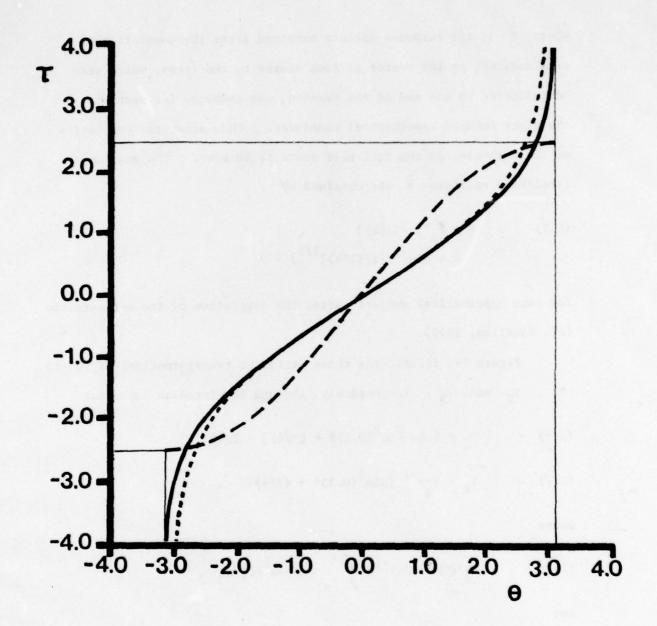


FIGURE 2-1

Transformations of  $\theta$  to  $\tau$  Which Result from Transforming the Constant Information Model with  $a_g$  = 0.25 and  $b_g$  = 0.00 As the Parameters, to the Linear Model with  $a_g$  = -2.50 and  $b_g$  = 2.50 (Broken Curve), to the Normal Ogive Model with  $a_g$  = 1.00 and  $b_g$  = 0.00 (Solid Curve), and to the Logistic Model with D = 1.70,  $a_g$  = 1.00 and  $b_g$  = 0.00 (Dotted Curve), Respectively.

(2.10) 
$$P_g^*(\tau_1) = (\tau_1 - \alpha_g)(\beta_g - \alpha_g)^{-1}$$
 for  $\alpha_g < \tau_1 < \beta_g$ ,

with  $\alpha_g = -2.50$  and  $\beta_g = 2.50$ , 2) the normal ogive model such that

(2.11) 
$$P_{g}^{**}(\tau_{2}) = (2\pi)^{-1/2} \int_{-\infty}^{a_{g}^{*}(\tau_{2}^{-b_{g}^{*})}} \exp[-u^{2}/2] du$$
 for  $-\infty < \tau_{2} < \infty$ 

with  $a_g^* = 1.00$  and  $b_g^* = 0.00$ , and 3) the logistic model such that

(2.12) 
$$P_{g}^{***}(\tau_{3}) = \left[1 + \exp\{-1.7a_{g}^{**}(\tau_{3}^{-b_{g}^{**}})\}\right]^{-1}$$
 for  $-\infty < \tau_{3} < \infty$ ,

with  $a_g^{\star\star}=1.00$  and  $b_g^{\star\star}=0.00$ , respectively. Table 2-1 presents the corresponding values of  $\tau_1$ ,  $\tau_2$  and  $\tau_3$  to each of the eight levels of  $\theta$ , at which hypothetical examinees are located.

#### TABLE 2-1

Values of  $\tau_1$ ,  $\tau_2$  and  $\tau_3$  Corresponding to the Eight Levels of  $\theta$ , Which Result from the Transformations from the Constant Information Model with  $a_g$  = 0.25 and  $b_g$  = 0.00 As the Parameters, to the Linear Model with  $\alpha_g$  = -2.50 and  $\beta_g$  = 2.50, to the Normal Ogive Model with  $a_g$  = 1.00 and  $a_g$  = 1.00 and  $a_g$  = 0.00 and to the Logistic Model with D = 1.70, a  $a_g$  = 1.00 and  $a_g$  = 0.00, Respectively.

Constant Information	Linear	Normal Ogive	Logistic		
θ	τ <sub>1</sub>	τ <sub>2</sub>	τ <sub>3</sub>		
-3.0	-2.494	-3.023	-3.930		
-2.2	-2.228	-1.604	-1.680		
-1.4	-1.611	-0.923	-0.900		
-0.6	-0.739	-0.379	-0.358		
0.2	0.250	0.125	0.118		
1.0	1.199	0.642	0.614		
1.8	1.958	1.235	1.240		
2.6	2.409	2.092	2.345		

#### III Results

Table 3-1 presents the sample mean,  $m_{\widehat{\theta}}$ , of the one hundred maximum likelihood estimates for each of the eight groups of examinees, which was obtained after the completion of each of the twenty sessions. It is observed that, as the number of items increases,  $m_{\widehat{\theta}}$  tends to approach the true ability  $\theta$  on each ability level. And yet we find substantial differences in the speed of convergence across the different groups. In this table, all the entries whose absolute discrepancies from the true  $\theta$  are greater than or equal to 0.15 are marked with \*\*\*, those the absolute discrepancies of which are greater than or equal to 0.10 and less than 0.15 are marked with \*\*\*, and those whose absolute discrepancies are greater than or equal to 0.05 and less than 0.05 are marked with \*.

We can see from Table 3-1 that, for Group 1, whose true ability level (= -3.0) is very close to the lower endpoint  $\theta$  (= - $\pi$ ), even after 200 items were administered, the discrepancy of  $m_{\hat{\theta}}$  from  $\theta$  is as large as 0.061. In contrast to this, for Group 5, whose ability level is 0.2, i.e., closest of all the eight levels to the midpoint (= 0.0) of the interval, after the completion of the first session, or the administration of only ten items, the discrepancy of  $m_{\hat{\theta}}$  from  $\theta$  is as small as 0.020. For the other six groups, with some fluctuations, there is a luminous tendency that, as the true ability level departs from the closer end point of the interval,  $\theta$  or  $\bar{\theta}$  (= $\pi$ ), the convergence of  $m_{\hat{\theta}}$  to  $\theta$  becomes faster.

Table 3-2 provides us with the sample standard deviation,  $s_{\hat{h}}$  ,

TABLE 3-1

Sample Mean of the Maximum Likelihood Estimates  $\hat{\theta}$  of the 100 Hypothetical Examinees of Each of the Eight Ability Levels, after Completing Each of the Twenty Sessions of Testing.

Group	1	2	3	4	5	6	7	8
Session	-3.0	-2.2	-1.4	-0.6	0.2	1.0	1.8	2.6
1	-3.142	-2.587	-1.453	-0.677	0.219	1.121	2.068	2.815
2	-3.142	-2.440		-0.637	0.214	1.082	1.855	2.816
3	-3.142	-2.342	-1.451	-0.625	0.195	1.045	1.811	2.759
4	-3.142	-2.302	-1.445	-0.596	0.200	1.042	1.811	2.731
5	-3.136	-2.26\$	-1.440	-0.566	0.203	1.046	1.810	2.732
6	-3.116	-2.264	-1.414	-0.559	0.220	1.037	1.789	2.706
7	-3.113	-2.241	-1.412	-0.571	0.205	1.024	1.795	2.702
8	-3.115	-2.233	-1.416	-0.563	0.202	1.025	1.794	2.696
9	-3.117	-2.227	-1.416	-0.577	0.201	1.025	1.806	2.686
10	-3.114	-2.232	-1.415	-0.580	0.197	1.024	1.796	2.670
11	-3.108	-2.220	-1.411	-0.583	0.196	1.017	1.801	2.674
12	-3.105	-2.211	-1.419	-0.587	0.186	1.019	1.799	2.668
13	-3.102	-2.218	-1.419	-0.598	0.188	1.018	1.803	2.660
14	-3.097	-2.212	-1.422	-0.594	0.190	1.017	1.803	2.649
15	-3.088	-2.209	-1.419	-0.596	0.196	1.015	1.801	2.635
16	-3.077	-2.198	-1.414	-0.598	0.198	1.013	1.802	2.639
17	-3.075	-2.197	-1.407	-0.599	0.199	1.016	1.801	2.629
18	-3.073	-2.191	-1.404	-0.608	0.202	1.017	1.804	2.615
19	-3.068	-2.192	-1.400	-0.612	0.206	1.020	1.804	2.613
20	-3.061	-2.194	-1.401	-0.611	0.206	1.020	1.800	2.615

\* 0.05 
$$\leq |m_{\hat{\theta}} - \theta| < 0.10$$

\*\* 0.10 
$$< |m_{\hat{\theta}} - \theta| < 0.15$$

\*\*\* 0.15 
$$\leq |m_{\hat{\theta}} - \theta|$$

TABLE 3-2

Sample Standard Deviation of the Maximum Likelihood Estimates  $\hat{\theta}$  of the 100 Hypothetical Examinees of Each of the Eight Ability Levels, after Completing Each of the Twenty Sessions of Testing.

Group	1	2	3	4	5	6	7	8	1/2
Session $\theta$	-3.0	-2.2	-1.4	-0.6	0.2	1.0	1.8	2.6	[I(θ)] <sup>-1/2</sup>
1	0.000	0.733	0.810	0.812*	0.68	0.822*	0.908*	0.570	0.632
2	0.000	0.556	0.490	0.464	0.460	0.517	0.592	0.461	0.447
3	0.000	0.465	0.415	0.359	0.403	0.368	0.482	0.410	0.365
4	0.000	0.407	0.330	0.319	0.328	0.297	0.383	0.374	0.316
5	0.057	0.297	0.285	0.273	0.280	0.274	0.321	0.347	0.283
6	0.113	0.270	0.261	0.267	0.252	0.235	0.267	0.310	0.258
7	0.114	0.226	0.232	0.251	0.229	0.216	0.248	0.286	0.239
8	0.107	0.212	0.225	0.235	0.219	0.211	0.222	0.276	0.224
9	0.100	0.214	0.222	0.217	0.203	0.197	0.216	0.262	0.211
10	0.102	0.214	0.216	0.205	0.198	0.196	0.203	0.239	0.200
11	0.109	0.200	0.202	0.213	0.187	0.187	0.184	0.233	0.191
12	0.110	0.186	0.191	0.198	0.174	0.178	0.175	0.223	0.183
13	0.115	0.180	0.178	0.188	0.162	0.166	0.173	0.210	0.175
14	0.118	0.171	0.167	0.173	0.163	0.162	0.169	0.208	0.169
15	0.124	0.165	0.166	0.172	0.154	0.157	0.170	0.189	0.163
16	0.130	0.163	0.163	0.164	0.146	0.157	0.163	0.187	0.158
17	0.131	0.159	0.160	0.157	0.140	0.150	0.158	0.169	0.153
18	0.132	0.152	0.157	0.153	0.137	0.146	0.154	0.143	0.149
19	0.135	0.149	0.157	0.153	0.135	0.142	0.156	0.138	0.145
20	0.135	0.141	0.153	0.155	0.130	0.137	0.150	0.121	0.141

\* 
$$0.05 \le |\mathbf{s}_{\hat{\theta}} - \{\mathbf{I}(\theta)\}^{-1/2}| < 0.10$$
  
\*\*  $0.10 \le |\mathbf{s}_{\hat{\theta}} - \{\mathbf{I}(\theta)\}^{-1/2}| < 0.15$   
\*\*\*  $0.15 \le |\mathbf{s}_{\hat{\theta}} - \{\mathbf{I}(\theta)\}^{-1/2}|$ 

\*\* 
$$0.10 \le |\mathbf{s}_{\hat{\mathbf{a}}} - \{\mathbf{I}(\theta)\}^{-1/2}| < 0.15$$

\*\*\* 
$$0.15 \le |s_{\hat{\theta}} - \{I(\theta)\}^{-1/2}|$$

which was calculated by using 10k as the denominator after the completion of the k-th session, of the one hundred maximum likelihood estimates for each group after the completion of each session. In the same table, the values of  $\left[I(\theta)\right]^{-1/2}$  are also given in the rightest hand column, for the purpose of comparison. Some of the entries of this table are marked with \*\*\*, \*\* or \*, following the same rule as we used for the sample mean  $m_{\hat{\theta}}$ , in accordance with the absolute discrepancy of the  $s_{\hat{\theta}}$  from  $\left[I(\theta)\right]^{-1/2}$ . We observe a tendency in Table 3-2, which is very similar to the one we have found in Table 3-1, the sample means.

We notice that for Group 1, whose ability level is very close to the lower endpoint of the interval, the sample standard deviation  $\mathbf{s}_{\hat{\theta}}$  equals zero after 10, 20, 30 and 40 items have been administered, respectively. This owes to the fact that the response patterns of all the one hundred examinees of this group consist, uniformly, of  $\mathbf{u}_{\mathbf{g}} = 0$ , up to the fourth session, and the likelihood function has, therefore, the terminal maximum at  $\theta = -\pi$ . Since we have from (1.2)

(3.1) 
$$P_{g}(-3.0) = \sin^{2}[(-3.0 + \pi)/4]$$
$$= 0.00125,$$

which is very close to zero, the above results are understandable.

Figure 3-1 presents the cumulative frequency ratios of the maximum likelihood estimates of the one hundred examinees of Group 1 (solid line), along with the normal distribution function,  $N(\theta, I(\theta)^{-1})$  (solid curve), after the completion of each of the twenty sessions.

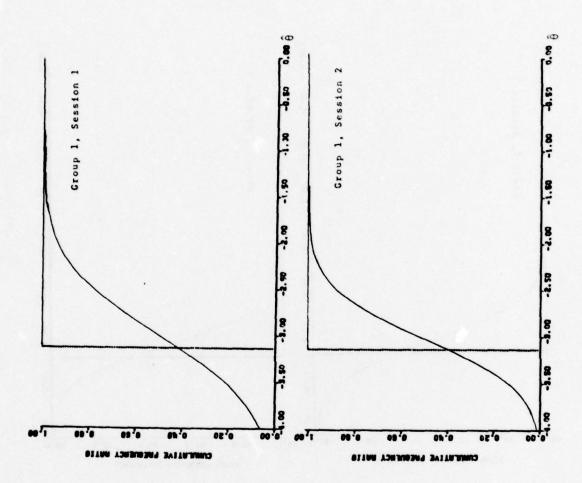


FIGURE 3-1

Cumulative Frequency Ratio of the Maximum Likelihood Estimates 6 of the 100 Hypothetical Examinees (Solid Line), with the Asymptotic Normal Distribution N(8, I(8)<sup>-1</sup>) (Solid Curve), and the Normal Distribution Function with the Sample Mean and Variance As Its Two Parameters (Dotted Curve), After Completing Each of the Twenty Sessions:

 $\theta = -3.0$ 

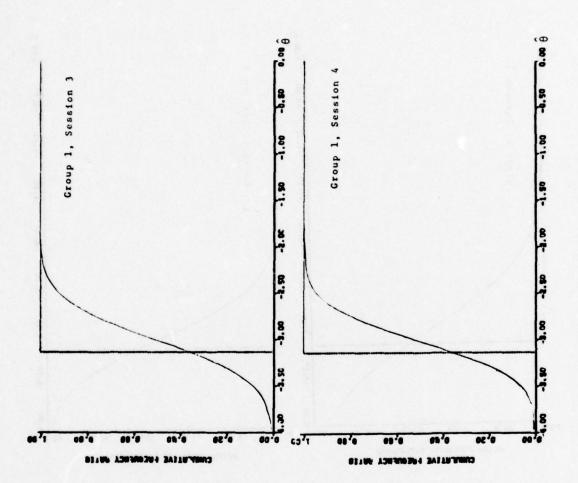


FIGURE 3-1 (Continued)

FIGURE 3-1 (Continued)

FIGURE 3-1 (Continued)

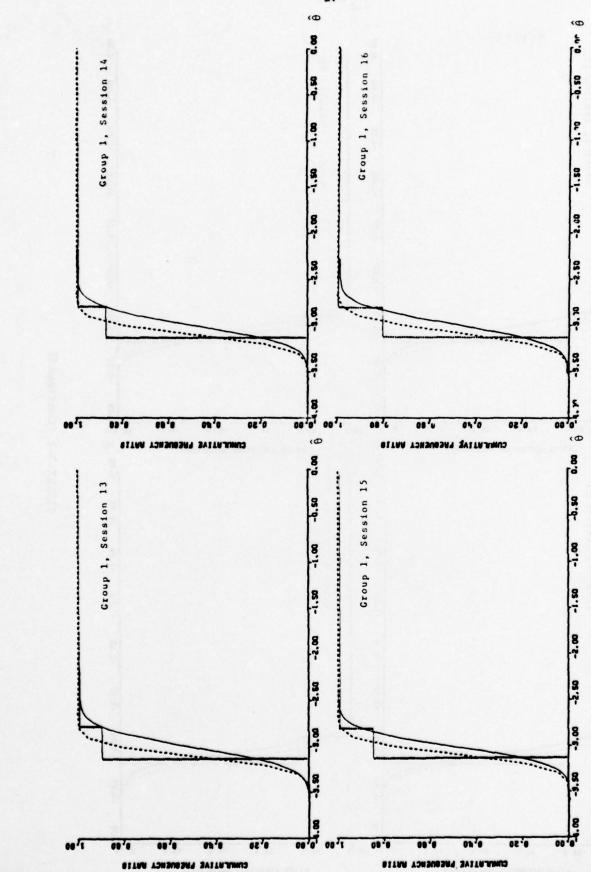


FIGURE 3-1 (Continued)

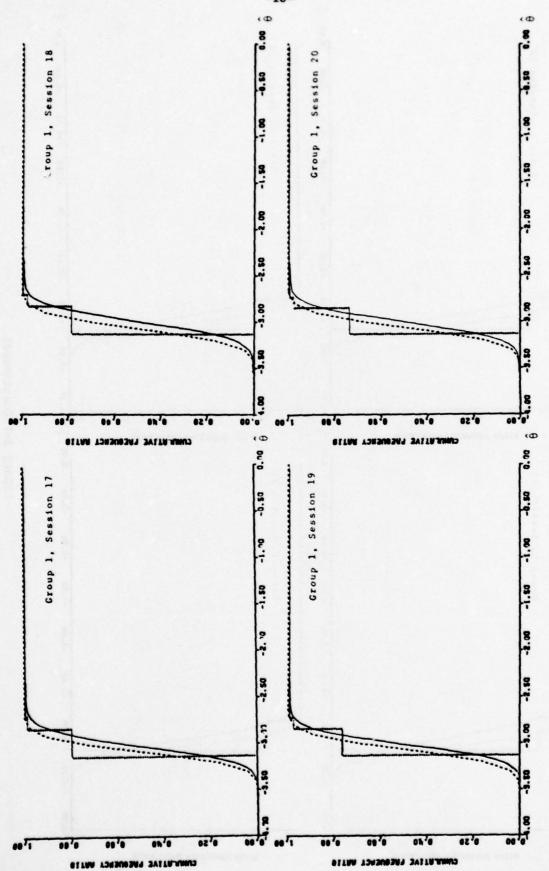


FIGURE 3-1 (Continued)

In the same figure, also presented are the corresponding normal distribution functions with  $m_{\widehat{A}}$  and  $s_{\widehat{A}}^2$  as the two parameters (dotted curves). These normal distribution functions coincide with the corresponding cumulative frequency ratios in the first four graphs of Figure 3-1, i.e., for sessions 1, 2, 3 and 4, since the sample variance is zero in these four cases. The discrepancy of the cumulative frequency ratio, and that of  $N(m_{\Omega}, s_{\Omega}^2)$ , from the other normal distribution function,  $N(\theta, I(\theta)^{-1})$ , are substantially large even in the last graph of Figure 3-1, where the maximum likelihood estimates were obtained on the basis of the It is also worth noting that the responses to two hundred items. cumulative frequency ratio of each of the twenty graphs of Figure 3-1 indicates a J-shape curve for the frequency distribution of the one hundred maximum likelihood estimates, the result which confirms our theoretical anticipation made earlier (cf. Samejima, 1979).

A similar set of twenty graphs as Figure 3-1 was made for Group 8, whose ability level (= 2.6) is the next closest to one of the endpoints of the interval,  $(-\pi, \pi)$ , and is presented as Figure 3-2. We find in this figure that, unlike the case of Group 1, the cumulative frequency ratio of the one hundred maximum likelihood estimates shows a good convergence to the normal distribution function,  $N(\theta, I(\theta)^{-1})$ , in each of the last three or four graphs, or after 170 or more items have been administered to the examinees. For this ability level, although a J-shape curve for the frequency distribution of the one hundred maximum likelihood estimates is indicated by the cumulative

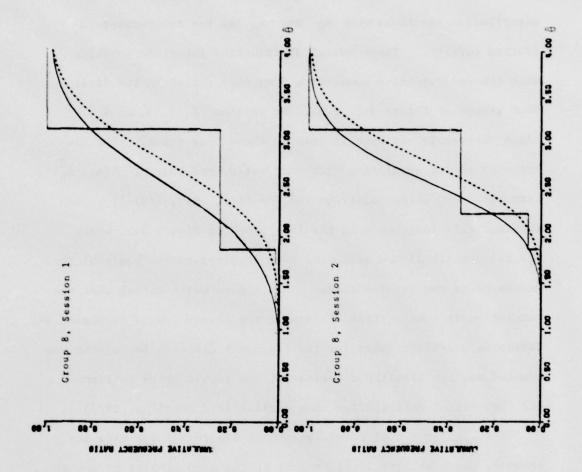


FIGURE 3-2

Cumulative Frequency Ratio of the Maximum Likelihood Estimates  $\theta$  of the 100 Hypothetical Examinees (Solid Line), with the Asymptotic Normal Distribution N( $\theta$ , I( $\theta$ )<sup>-1</sup>) (Solid Curve), and the Normal Distribution Function with the Sample Mean and Variance As Its Two Parameters (Dotted Curve), After Completing Each of the Twenty Sessions:

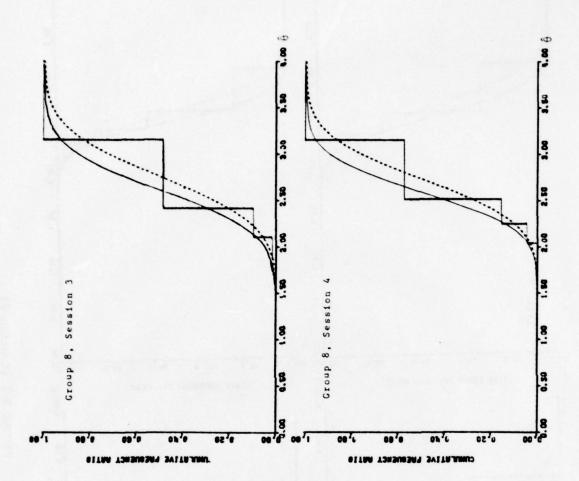


FIGURE 3-2 (Continued)

CUMULATIVE PREQUENCY RATIO

UMALATIVE PACOUCACT ANTI-

FIGURE 3-2 (Continued)

FIGURE 3-2 (Continued)

FIGURE 3-2 (Continued)

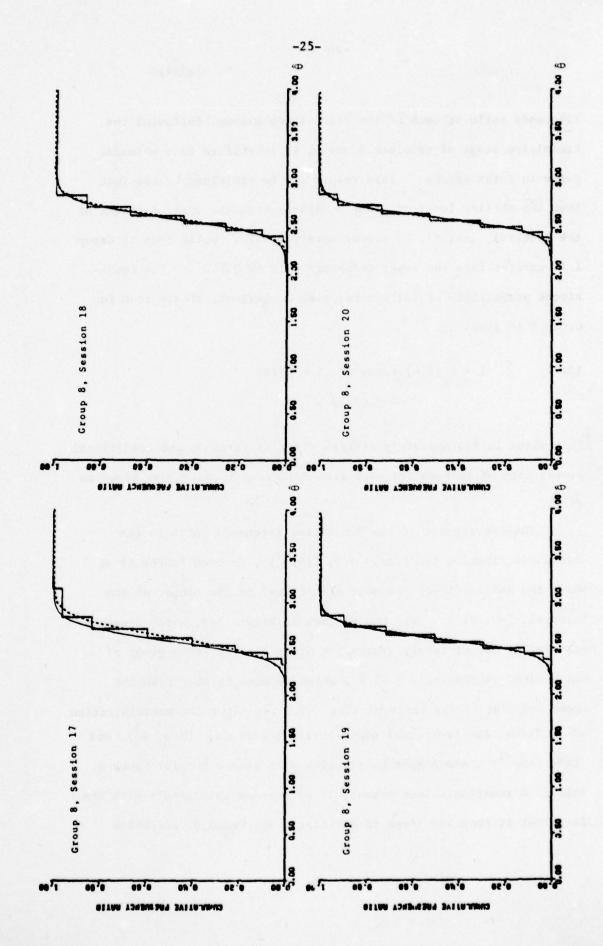


FIGURE 3-2 (Continued)

frequency ratio of each of the first three graphs, following the transition stage of sessions 4 and 5, it is shifted to a unimodal curve in later graphs. This result can be explained by the fact that the ability level of Group 8 departs from the upper endpoint of the interval,  $(-\pi, \pi)$ , by approximately 0.542, while that of Group 1 is greater than the lower endpoints only by 0.142. The conditional probability of failure for each equivalent, binary item for Group 8 is given by

(3.2) 
$$1 - P_g(2.6) = \cos^2[(2.6 + \pi)/4]$$

$$= 0.01822.$$

This value is approximately fifteen times as large as the conditional probability of success for each item for Group 1, which was given as (3.1).

The convergence of the cumulative frequency ratio to the normal distribution function,  $N(\theta, I(\theta)^{-1})$ , is much faster if we shift the ability level one more step closer to the center of the interval,  $(-\pi, \pi)$ . This can be seen in Figure 3-3, which presents the similar set of twenty graphs for Group 2. For this group of one hundred examinees,  $\theta$  = -2.2, which is much farther from the lower endpoint of the interval than -3.0. After the administration of 70 items, the two normal distribution functions,  $N(m_{\widehat{\theta}}, s_{\widehat{\theta}}^2)$  and  $N(\theta, I(\theta)^{-1})$ , are almost overlapping with each other for Group 2. This is a remarkable improvement, if we compare this result with the fact that it took 170 items to administer for Group 8, for which

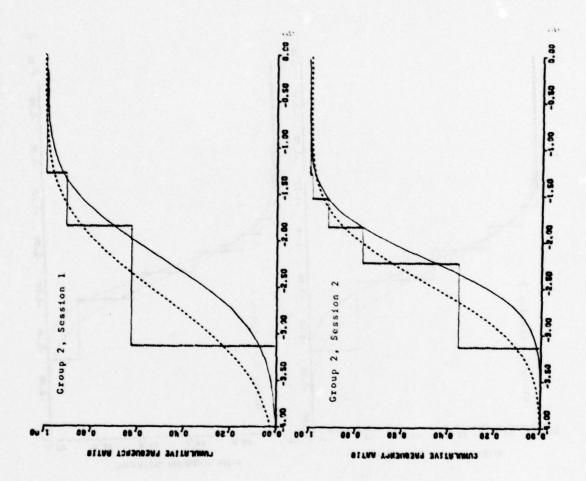


FIGURE 3-3

Cumulative Frequency Ratio of the Maximum Likelihood Estimates  $\theta$  of the 100 Hypothetical Examinees (Solid Line), with the Asymptotic Normal Distribution  $N(\theta, I(\theta)^{-1})$  (Solid Curve), and the Normal Distribution Function with the Sample Mean and Variance As Its Two Parameters (Dotted Curve), After Completing Each of the Twenty Sessions:

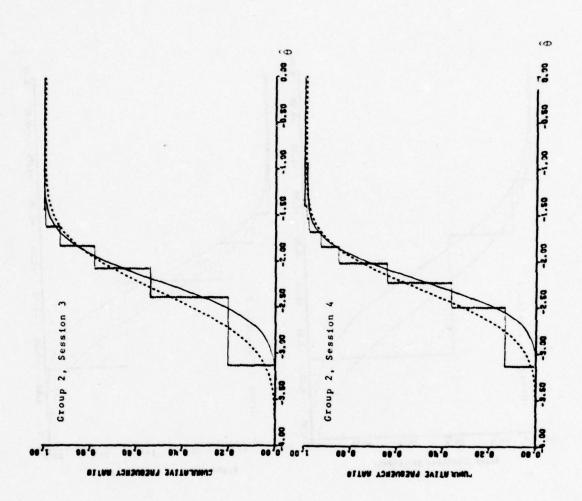


FIGURE 3-3 (Continued)

FIGURE 3-3 (Continued)

FIGURE 3-3 (Continued)

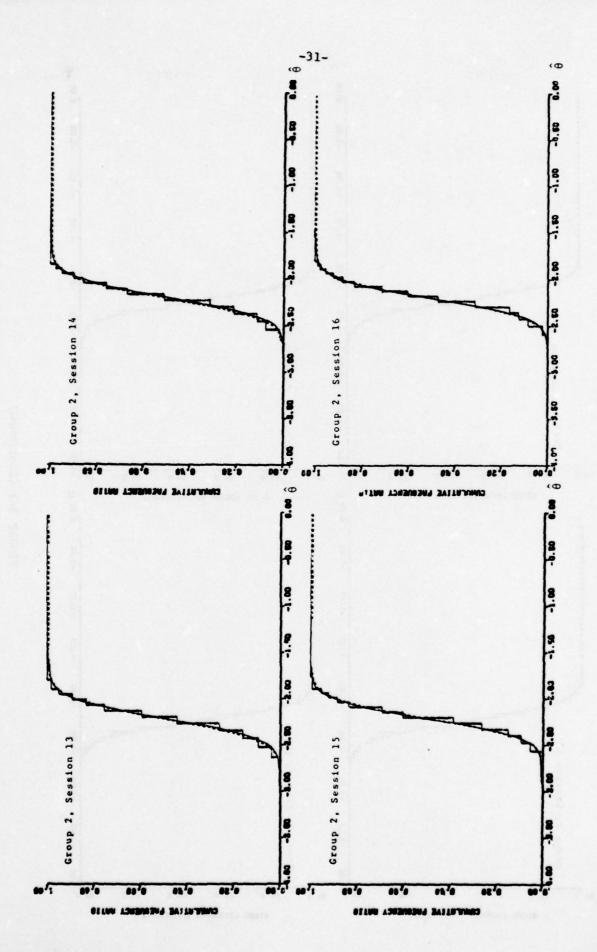


FIGURE 3-3 (Continued)

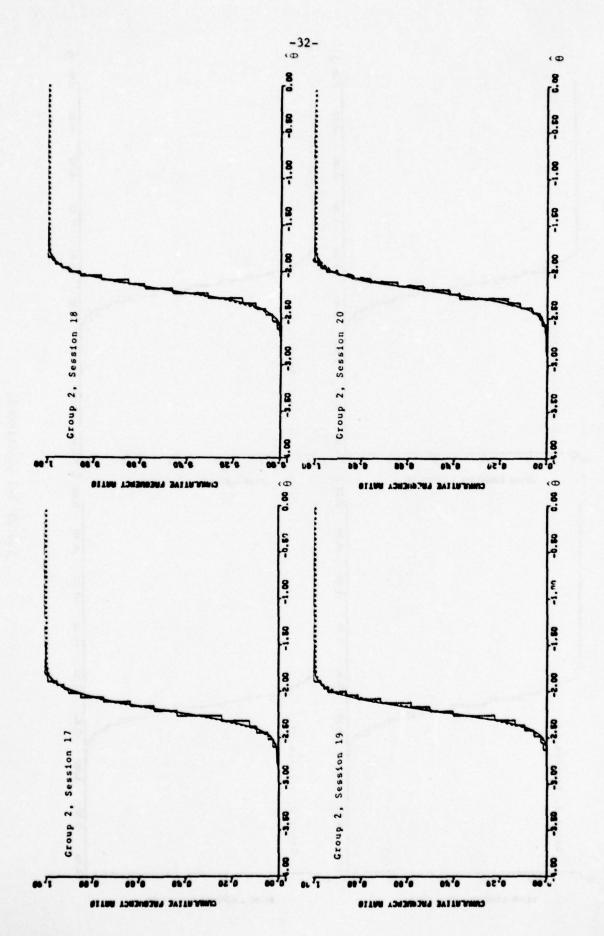


FIGURE 3-3 (Continued)

 $\theta$  = 2.6, to achieve this level of closeness. A J-shape frequency distribution of the maximum likelihood estimate is indicated by the cumulative frequency ratio only in the first graph of Figure 3-3, i.e., after ten items have been administered to the examinees, and, after that, unimodal frequency distributions are observed.

The remarkable improvement in the speed of convergence of the cumulative frequency ratio of one hundred maximum likelihood estimates for Group 2 in comparison with Group 8 can be explained from the difference between the conditional probability of success for Group 2 and that of failure for Group 8. The former conditional probability is given by

(3.3) 
$$P_{g}(-2.2) = \sin^{2}[(-2.2 + \pi)/4]$$
$$= 0.05440$$

whereas the latter conditional probability was already obtained in (3.2), which is approximately one-third of  $P_g(-2.2)$ . This difference between the two conditional probabilities indicates that the probability with which the examinee of Group 2 obtains the strictly decreasing likelihood function with the terminal maximum at  $\theta=-\pi$  is much smaller than the one with which the examinee of Group 8 obtains the strictly increasing likelihood function with the terminal maximum at  $\theta=\pi$ , especially when the number of items is substantially large.

So far we have observed the results of the three groups of one hundred examinees whose ability levels are close to one of the

endpoints of the interval,  $(-\pi, \pi)$ . In contrast to those results, we shall observe a similar set of twenty graphs for Group 5, for which  $\theta$  = 0.2, i.e., closest to the center of the interval. Figure 3-4 presents these results.

We can see in this figure that, even in the first graph, or after the administration of only ten items, the two normal distribution functions,  $N(m_{\widehat{\theta}}, s_{\widehat{\theta}}^2)$  and  $N(\theta, I(\theta)^{-1})$ , are very close. We also notice that there are more varieties of different values of the maximum likelihood estimate, and thus the cumulative frequency ratio is much smoother, in earlier graphs, in comparison with those for the other three groups we observed earlier. There is no doubt that the speed of convergence of the distribution of the maximum likelihood estimate to the normality is much higher on this level of  $\theta$ .

Similar sets of twenty graphs for the remaining four groups,
Groups 3, 4, 6 and 7, are presented in Appendix I as Figures A-1-1,
A-1-2, A-1-3 and A-1-4, respectively. We can say from these results
that, except for Group 7, whose ability level is 1.8, the speed of
convergence of the cumulative frequency ratio of the maximum likelihood
estimate is quite high. The conditional probability of success
in one binary item for each of these four groups of subjects is given
in Table 4-4 which will be presented in the next section.

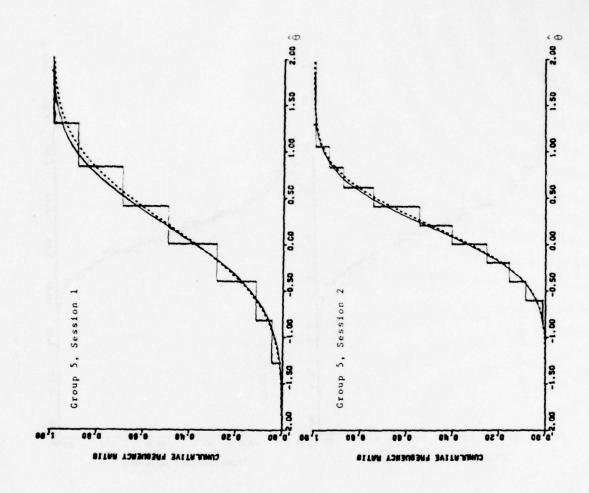


FIGURE 3-4

Cumulative Frequency Ratio of the Maximum Likelihood Estimates <sup>6</sup> of the 100 Hypothetical Examinees (Solid Line), with the Asymptotic Normal Distribution N(<sup>6</sup>, I(<sup>6</sup>)<sup>-1</sup>) (Solid Curve), and the Normal Distribution Function with the Sample Mean and Variance As Its Two Parameters (Dotted Curve), After Completing Each of the Twenty Sessions:

 $\theta = 0.2$ 

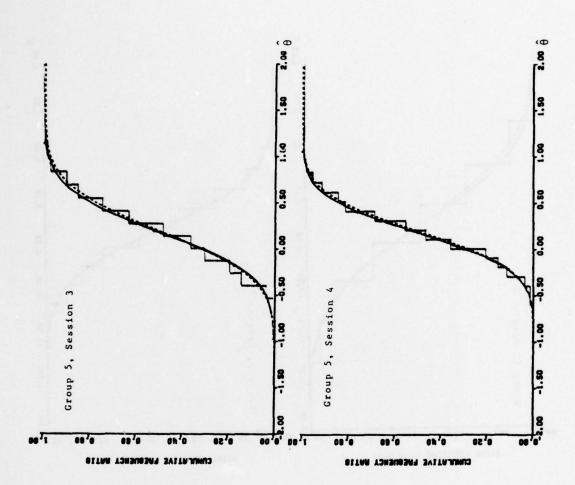


FIGURE 3-4 (Continued)

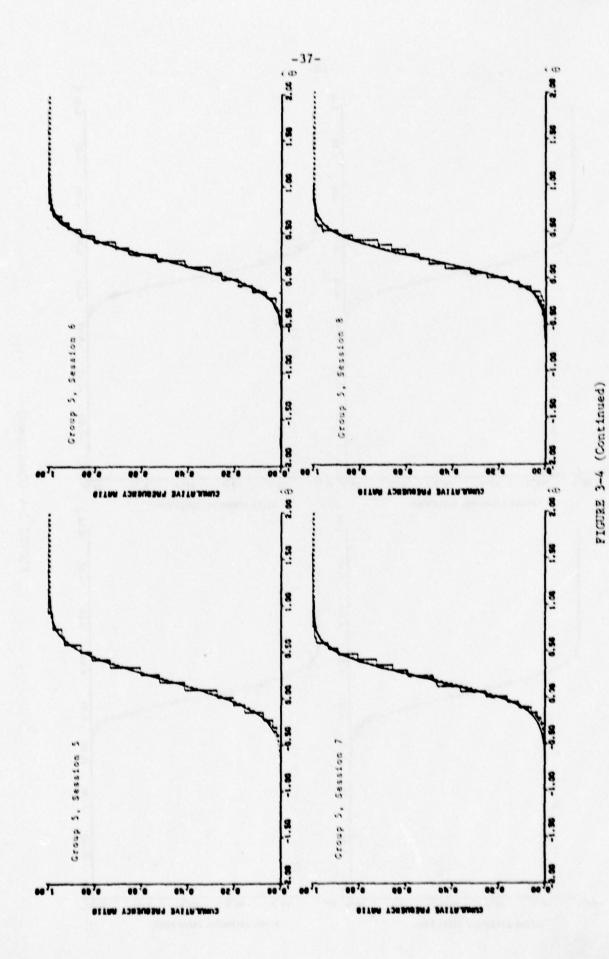


FIGURE 3-4 (Continued)

FIGURE 3-4 (Continued)

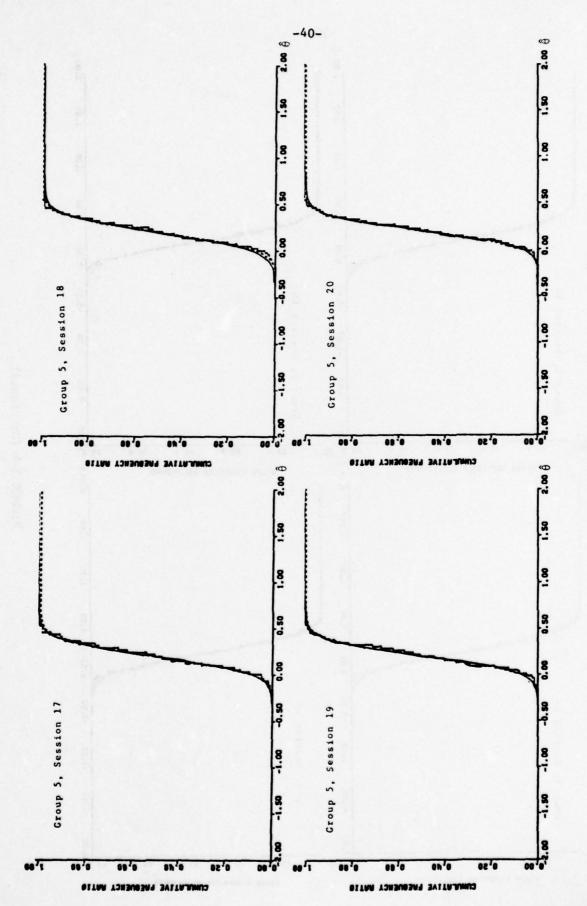


FIGURE 3-4 (Continued)

## IV Further Observation of the Results

We have observed the sample mean,  $m_{\widehat{\theta}}$ , and the sample standard deviation,  $s_{\widehat{\theta}}$ , of the one hundred maximum likelihood estimates of each of the eight groups of examinees, after the completion of each of the twenty sessions of testing. We have also observed the cumulative frequency ratio in each case, in comparison with the normal distribution function,  $N(\theta,\ I(\theta)^{-1})$ . For further observation, here we consider the skewness index,  $c_{1\widehat{\theta}}$ , which is defined by

$$c_{1\hat{\theta}} = m_{3\hat{\theta}} s_{\hat{\theta}}^{-3} ,$$

where  $m_{3\hat{H}}$  is the third sample moment about the mean, given by

(4.2) 
$$\mathfrak{m}_{3\hat{\theta}} = (1/100) \sum_{i=1}^{100} (\hat{\theta}_i - \mathfrak{m}_{\hat{\theta}})^3$$
.

Table 4-1 presents the value of this sample skewness index,  $c_{1\widehat{\theta}}$ , for the set of the one hundred maximum likelihood estimates of each of the eight groups of examinees, obtained after the completion of each of the twenty sessions. From (4.1), it is obvious that a positive value in the table indicates a positive skewness, and a negative value means a negative skewness. In this table, again the three marks, \*\*\*, \*\* and \*, are assigned to some of the entries. The first mark is attached to all the values of  $c_{1\widehat{\theta}}$ , whose absolute values are greater than or equal to 1.5, the second one is for those the absolute values of which are greater than or equal to 1.0 and less than 1.5, and the third one is assigned to those whose absolute values are greater than or equal to 0.5 and less than

TABLE 4-1

Sample Skewness Index of the Maximum Likelihood Estimates  $\;\hat{\theta}\;$  of the 100Hypothetical Examinees of Each of the Eight Ability Levels, after Completing Each of the Twenty Sessions of Testing.

Group	1	2	3	4	5	6	7	8
Session	-3.0	-2.2	-1.4	-0.6	0.2	1.0	1.8	2.6
1	10 <u></u> 1	0.770	-0.782	-0.622	-0.120	0.668	-0.071	-1.196
2		-0.134	-0.272	-0.367	-0.052	0.599	0.735	-0.786
3		-0.544	-0.720	0.071	0.071	0.058	0.829	-0.267
4		-0.676	-0.431	0.317	0.130	-0.111	0.771	-0.033
5	9.849*	-0.626	-0.545	0.241	0.225	-0.105	0.762	0.110
6	4.129	-0.339	-0.526	0.090	0.070	-0.007	-0.046	0.344
7	3.705	0.050	-0.502	0.078	0.074	-0.376	0.022	0.506
8	3.705	-0.245	-0.451	0.156	0.141	-0.275	0.026	0.497
9	3.705	-0.371	-0.382	0.218	0.349	-0.328	0.064	0.609
10	3.371*	-0.277	-0.345	0.140	0.597	-0.311	-0.033	0.717
11	2.865	-0.237	-0.382	0.222	0.459	-0.132	-0.177	0.664
12	2.667	-0.167	-0.249	0.433	0.421	-0.022	0.022	0.709
13	2.577*	-0.183	-0.367	0.246	0.389	-0.020	0.066	0.801
14	2.271*	-0.214	-0.406	0.198	0.242	-0.029	0.182	0.851
15	1.912*	-0.122	-0.355	0.252	0.212	-0.115	0.266	0.862
16	1.545*	-0.157	-0.293	0.255	0.152	-0.064	0.308	0.899
17	1.500	-0.121	-0.407	0.015	0.068	-0.324	0.154	0.938
18	1.448	-0.183	-0.490	0.048	0.027	-0.232	0.228	0.587
19	1.380	-0.316	-0.445	0.193	0.055	-0.048	0.172	0.697
20	1.189	-0.271	-0.443	0.187	0.067	-0.032	0.200	0.287

\* 
$$0.5 \le |c_{1\hat{\theta}}| < 1.0$$
\*\*  $1.0 \le |c_{1\hat{\theta}}| < 1.5$ 
\*\*\*  $1.5 \le |c_{1\hat{\theta}}|$ 

1.0. We notice that the first four entries of the column of Group

1 in the table are indeterminate, since both the third and second

moments about the mean are zero in each of these four cases.

It is interesting to note that, for Group 2, the sign of the index is reversed between the first and second sessions, and, for Group 8, it is reversed between the fourth and fifth sessions, while it is consistently positive for Group 1. Unlike Tables 3-1 and 3-2, which were presented in the preceding section and are for the sample mean and the sample standard deviation, respectively, the configuration of the values of the sample skewness index in Table 4-1 is somewhat more complicated and seemingly more difficult to interpret.

This problem of interpretation will be solved, however, if we pay attention to the frequencies of the two extreme values of the maximum likelihood estimate,  $-\pi$  and  $\pi$ . Table 4-2 presents the frequency distribution of these two values for each group and for each session. In this table, the numbers appearing in the first four columns are the frequencies of  $-\pi$ , and those in the last three columns are those of  $\pi$ . For Group 5, the frequencies are zero for both  $-\pi$  and  $\pi$  throughout the twenty sessions.

We can see in Table 4-2 that, for Group 1, even after the administration of 200 items, the terminal maximum for the likelihood function at the lower end of the interval  $(-\pi, \pi)$  occurred 73 times out of 100. This explains the J-shape cumulative frequency ratios of the maximum likelihood estimate for Group 1 throughout the

TABLE 4-2

Frequencies of  $-\pi$  As the Maximum Likelihood Estimate Appearing in the First 4 Columns and Those of  $\pi$  Appearing in the Last 3 Columns. The Total Number of Maximum Likelihood Estimates in Each Cell is 100.

Group	1	2	3	4	5	6	7	8
Session $\theta$	-3.0	-2.2	-1.4	-0.6	0.2	1.0	1.8	2.6
1	100	62	13	3	0	8	37	75
2	100	35	1	0	0	1	11	66
3	100	20	1	0	0	0	5	52
4	100	13	0	0	0	0	2	43
5	99	4	0	0	0	0	1	39
6	95	2	0	0	0	0	0	30
7	94	0	0	0	0	0	0	26
8	94	0	0	0	0	0	0	23
9	94	0	0	0	0	0	0	20
10	93	0	0	0	0	0	0	15
11	91	0	0	0	0	0	0	14
12	90	0	0	0	0	0	0	12
13	89	0	0	0	0	0	0	10
14	87	0	0	0	0	0	0	9
15	84	0	0	0	0	0	0	6
16	80	0	0	0	0	0	0	6
17	79	0	0	0	0	0	0	4
18	78	0	0	0	0	0	0	1
19	76	0	0	0	0	0	0	1
20	73	0	0	0	0	0	0	0

twenty sessions of testing, as we have observed in the preceding section. On the other hand, a close observation of the relationship between the skewness indices and the frequencies of the two extreme values of the maximum likelihood estimate reveals that, for Groups 8, 2 and 7, the reversal of the sign of the skewness index occurs somewhere around the frequency 40 of the extreme values of the maximum likelihood estimate. Why is it so?

To answer this question, we must call our attention to the fact that there is a substantial gap between one of the extreme values of the maximum likelihood estimate and the adjacent one. To be more specific, for Session 1, or after the administration of 10 items, the value of the maximum likelihood estimate corresponding to the second best test score, 9, is 1.8546, which is obtained through (2.5) with k=1. Thus the difference of this value of the maximum likelihood estimate from  $\pi$  is as large as 1.2870. The same amount of discrepancy exists between  $-\pi$  and the adjacent maximum likelihood estimate, -1.8546, which corresponds to the second lowest test score, 1. Even for Session 20, or after all the 200 items have been administered, this discrepancy is still as large as 0.2831, since the maximum likelihood estimate corresponding to the second highest test score, 199, is 2.8585, and the one corresponding to the second lowest test score, 1, is -2.8585.

We notice that, when  $-\pi$ , one of the two extreme values of the maximum likelihood estimate, has a high frequency, it makes the total frequency distribution of the maximum likelihood estimate

J-shaped, causing a positive skewness to it. On the other hand, when such an extreme value has a low frequency, because of the long distance from the adjacent value of the maximum likelihood estimate, it creates a long tail on the left hand side of the total frequency distribution, and, therefore, causes a negative skewness to the frequency distribution. We can easily see that the same logic applies for the frequency of the other extreme value,  $\pi$ , of the maximum likelihood estimate, and negative and positive skewnesses. Considering these facts, we are in a good position to understand the meaning of the total configuration of the skewness indices in Table 4-1. It should also be pointed out that those which involve neither of the two extreme values,  $-\pi$  and  $\pi$ , of the maximum likelihood estimate are small enough to be considered as sampling fluctuations. In fact, only a few of them are marked with \*, and none of the others exceeds 0.5 in absolute value.

If we combine the result in Table 4-1 with the sample means of the maximum likelihood estimate, which are given in Table 3-1 in the preceding section, the effect of the two extreme values of the maximum likelihood estimate is more conspicuous. We notice that all the sample means, which involve one or more  $-\pi$  or  $\pi$ , are shifted from the true values of  $\theta$  to the directions indicated by the signs of the extreme values. It is also noted that all the sample means, which involve neither  $-\pi$  nor  $\pi$ , have very small discrepancies from the true values of  $\theta$ , i.e., less than 0.05 in absolute value.

Table 4-3 presents the kurtosis index,  $\,c_{2\hat{\theta}}^{}$  , which is defined by

(4.3) 
$$c_{2\hat{\theta}} = m_{4\hat{\theta}} s_{\hat{\theta}}^{-4} - 3$$
,

where m<sub>40</sub> is the fourth sample moment of the one hundred maximum likelihood estimates, for which we can write

(4.4) 
$$m_{4\hat{\theta}} = (1/100) \sum_{i=1}^{100} (\hat{\theta}_i - m_{\hat{\theta}})^4 ,$$

for each of the eight groups of examinees after the completion of each of the twenty sessions. The three types of marks, \*\*\*, \*\* and \*, are assigned in this table by following the same rule as we have used for the skewness index in Table 4-1. As is well known, this kurtosis index is defined in comparison with the normal distribution, i.e., a positive value indicates a greater peakedness, and a negative value a lesser peakedness, than the normal distribution.

We notice that, again in these results, there exists a certain effect of the extreme values,  $-\pi$  and  $\pi$ , of the maximum likelihood estimate. In fact, we observe two different transitions there, one of which is from positive to negative values, and the other from negative to positive values. The former transition stage corresponds, approximately, to the frequency 75 of  $-\pi$  or  $\pi$ , and the latter transition stage corresponds, roughly, to the frequency 13. Sampling fluctuations appear to be greater for the kurtosis index than for the skewness index, the fact which coincides with

TABLE 4-3

Sample Kurtosis Index of the Maximum Likelihood Estimates  $\hat{\theta}$  of the 100 Hypothetical Examinees of Each of the Eight Ability Levels, after Completing Each of the Twenty Sessions of Testing.

Group	1	2	3	4	5	6	7	8
Session	-3.0	-2.2	-1.4	-0.6	0.2	1.0	1.8	2.6
1		-0.866	0.027	0.980	-0.499	0.567	-1.141	-0.472
2		-1.278	0.720	0.270	-0.525	1.576	0.080	-1.184
3		-0.617	1.429	-0.179	-0.699	0.267	0.888	-1.608
4		0.124	0.940	0.136	-0.296	0.084	1.291	-1.505
5	95.010	1.299	0.458	0.291	-0.308	-0.108	2.071	-1.520
6	15.053	0.930	0.310	-0.574	-0.484	-0.326	-0.421	-1.142
7	11.730	-0.423	0.622	-0.465	-0.817	-0.236	0.087	-0.973
8	11.730	-0.010	0.087	-0.303	-0.703	-0.394	0.225	-0.758
9	11.730	0.114	-0.477	-0.228	-0.192	-0.343	0.752	-0.552
10	9.361	0.003	-0.317	-0.279	0.423	0.170	1.404	-0.152
11	6.210	0.021	-0.402	0.115	-0.212	0.414	0.608	-0.085
12	5.111	-0.307	-0.505	0.242	-0.501	0.349	0.298	0.183
13	4.893	-0.117	-0.319	-0.052	-0.510	0.517	0.305	0.472
14	3.358	-0.164	-0.119	0.030	-0.622	1.405	0.121	0.492
15	1.804	0.034	-0.194	0.184	-0.767	1.264	-0.077	0.867
16	0.499	-0.080	-0.303	0.293	-0.696	0.659	0.014	0.895
17	0.424	0.027	-0.085	0.358	-0.746	0.699	0.170	1.278
18	0.301	0.027	0.167	0.535	-0.650	0.375	-0.015	0.451
19	0.284	0.069	-0.042	0.306	-0.695	-0.031	-0.007	0.858
20	-0.236	-0.117	-0.307	0.793	-0.483	0.276	0.092	-0.628

\* 
$$0.5 \leqslant |c_{2\hat{\theta}}| < 1.0$$

\*\* 1.0 
$$< |c_{2\hat{\theta}}| < 1.5$$

our general knowledge of greater sampling fluctuations for higher order moments.

Table 4-4 presents the conditional probability of success,  $P_g(\theta)$ , for a single item, which is given by (1.2), for each of the eight ability levels, together with the item information function  $I_g(\theta)$ , and three other item information functions obtainable after the transformation of  $\theta$  to  $\tau_1$ ,  $\tau_2$  and  $\tau_3$  through (2.6), (2.7) and (2.9), respectively. These item information functions are obtained by multiplying  $I_g(\theta)$  with the square of the derivative of  $\theta$  with respect to the corresponding transformed variables, respectively (cf. Samejima, 1979).

TABLE 4-4

Conditional Probability of Success,  $P_g(\theta)$ , in a Single Equivalent Item g, Item Information Function,  $I_g(\theta)$ , and Three Other Item Information Functions,  $I_g^*(\tau_1)$ ,  $I_g^{**}(\tau_2)$  and  $I_g^{***}(\tau_3)$ , Obtained After Transforming Ability  $\theta$  to  $\tau_1$ ,  $\tau_2$  and  $\tau_3$ , Respectively.

θ	Pg(0)	I <sub>g</sub> (0)	1*(τ <sub>1</sub> )	Ι**(τ <sub>2</sub> )	Ι***(τ <sub>3</sub> )
-3.0	0.001253	0.25	31.9762	0.0137	0.0036
-2.2	0.054396	0.25	0.7776	0.2364	0.1487
-1.4	0.177891	0.25	0.2735	0.4639	0.4227
-0.6	0.352240	0.25	0.1753	0.6041	0.6594
0.2	0.549917	0.25	0.1616	0.6330	0.7153
1.0	0.739713	0.25	0.2078	0.5471	0.5564
1.8	0.891663	0.25	0.4141	0.3581	0.2792
2.6	0.981779	0.25	2.2360	0.1118	0.0517

## V Discussion and Conclusions

We have found out that the convergence of the cumulative frequency distribution of the maximum likelihood estimate to the normal distribution function,  $N(\theta, I(\theta)^{-1})$ , is substantially slower for the levels of ability close to one of the two endpoints of the interval,  $(-\pi, \pi)$ , in comparison with the levels of ability close to the center of the interval. It has become obvious that the frequency of one of the two extreme values,  $-\pi$  and  $\pi$  , of the maximum likelihood estimate is a good indicator of the degree of convergence, since it affects, strongly, the moments of the frequency distribution. One important implication of this finding is that we should avoid using the normal distribution,  $N(\theta, I(\theta)^{-1})$ , as the approximation to the conditional distribution of the maximum likelihood estimate, unless there are a large enough number of items so that the conditional probability of the "all success" response pattern, and that of the "all failure" response pattern, are neglibly small.

To pursue this further, the conditional probability of the "all success" response pattern, and that of the "all failure" response pattern, were calculated from  $P_g(\theta)$ , which is presented for each of the eight different levels of ability in Table 4-4 of the preceding section. Since all the items are equivalent, these conditional probabilities are  $P_g(\theta)^n$  and  $\left[1-P_g(\theta)\right]^n$ , respectively. For  $\theta$  = -3.0, or the ability level of Group 1, when we increase the number of items to 256, 512, 1,024, 2,048 and 4,096, the conditional

probability of the "all failure" response pattern becomes 0.725, 0.526, 0.277, 0.077 and 0.006, respectively. This means that, even if more than 1,000 items are administered, we must expect approximately 28 examinees out of 100 will obtain the "all failure" response pattern, and we need more than 4,000 items to reduce it to 0.6 person out of 100 examinees. These outcomes are exactly the same for  $\theta$  = 3.0 , when we consider the conditional probability of the "all success" response pattern. We may conclude, therefore, that for the values of  $\theta$  3.0 or greater in absolute value we should give up, totally, the idea of using the normal approximation to the conditional distribution of the maximum likelihood estimate.  $\theta$  = 2.6, if we increase the number of items from 32 to 64, 128, 256 and 512, the conditional probabilities of the "all success" response pattern are 0.555, 0.308, 0.095, 0.009 and 0.00008, respectively. Thus we can say that, for this level of ability, if we use approximately 300 items, chances are slim for one person out of 100 examinees to obtain  $\pi$  as the maximum likelihood estimate. This number, 300, is still too large for practical purposes, however, and we must say that in the vicinity of  $\theta$  = 2.6, and also that of  $\theta$  = -2.6, the idea of using the normal approximation to the conditional distribution of the maximum likelihood estimate is unrealistic. situation is much more ameliorated, however, if we switch to the ability level,  $\theta = -2.2$ . On this level of ability, for 8, 16, 32, 64 and 128 items, the conditional probabilities of the "all failure" response pattern are 0.639, 0.409, 0.167, 0.028 and 0.0008,

respectively. Thus we can conclude that, as long as we have 100 items, it is unlikely to happen that one person out of one hundred examinees obtains - m as his maximum likelihood estimate. is well exemplified and supported by the result of our Monte Carlo study, which is given in Table 4-2 of the preceding section. other words, we can be assured that 100 items are enough as far as we use the subinterval of  $\theta$  , (-2.2, 2.2) , for which the normal approximation to the conditional distribution of the maximum likelihood estimate is adopted. This number of items will be reduced to half, i.e., 50 , if we switch the subinterval of  $\,\theta\,$  from (-2.2, 2.2) to (-1.8, 1.8), since for  $\theta = 1.8$  the conditional probabilities of the "all success" response pattern are 0.632, 0.400, 0.160, 0.025 and 0.0006, for 4, 8, 16, 32 and 64 items, respectively. If we need to use, say, only 35 items, we must narrow down the interval further to (-1.4, 1.4) , for the conditional probabilities of the "all failure" response pattern at  $\,\theta$  = -1.4 , or those of the "all success" response pattern at  $\theta = 1.4$ , are 0.676, 0.457, 0.209, 0.044 and 0.002 for 2, 4, 8, 16 and 32 items, respectively. notice that the gain in the reduction of the number of items is decreasing compared with the amount of sacrifice in the interval length, and it is probably meaningless to consider further reductions of items. If we wish to use a set of "unknown" equivalent items as a substitute for the Old Test, it is advisable to aim at including 35 to 50 equivalent items in the new item pool. This can be done through the content analysis of the items, and their proportions

correct in the preliminary study, etc. As an additional information, the conditional probabilities of the "all success" response pattern for the ability levels 1.0, 0.6 and 0.2, or those of the "all failure" response pattern for the ability levels -1.0, -0.6 and -0.2, are: 0.547, 0.299, 0.090 and 0.008; 0.420, 0.176, 0.031 and 0.001; and 0.302, 0.091, 0.008 and 0.00007; for 2, 4, 8 and 16 items, respectively.

The conclusion made in the preceding paragraph is supported by the result of our Monte Carlo study, which was made in earlier sections. In the graph of Session 10 in Figure 3-3, that of Session 5 in Figure A-1-4, and that of Session 4 in Figure A-1-1, which involve 100 items for  $\theta = -2.2$ , 50 items for  $\theta = 1.8$  and 40 items for  $\theta = -1.4$ , respectively, the normal distribution function,  $N(\theta, I(\theta)^{-1})$  and the cumulative frequency ratio are sufficiently We find in Tables 3-1, 3-2, 4-1 and 4-3 that the sample mean, the sample standard deviation, the sample skewness index and the sample kurtosis index are: -2.232 [  $\theta$  = -2.200 ], 0.214 [  $I(\theta)^{-1/2}$ = 0.200 ], -0.277 and 0.003 in the first case, 1.810 [  $\theta$  = 1.800 ], 0.321 [  $I(\theta)^{-1/2} = 0.283$  ], 0.762 and 2.071 in the second case, and -1.445 [  $\theta = -1.400$  ], 0.330 [  $I(\theta)^{-1/2} = 0.316$  ], -0.431 and 0.940 in the third case, respectively. We can say that, at least, the sample mean and the sample standard deviation are very close to the asymptotic values in each case.

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APPENDIX I

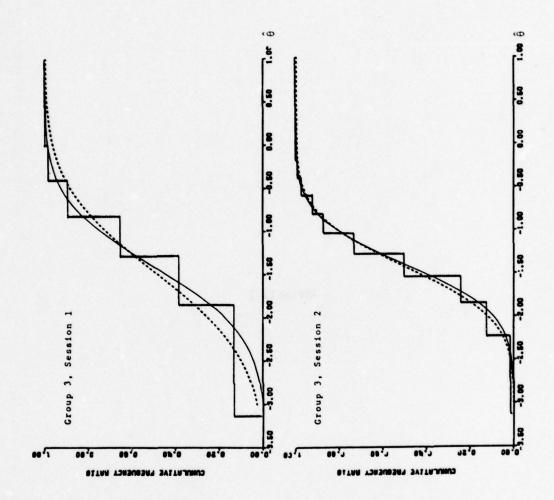
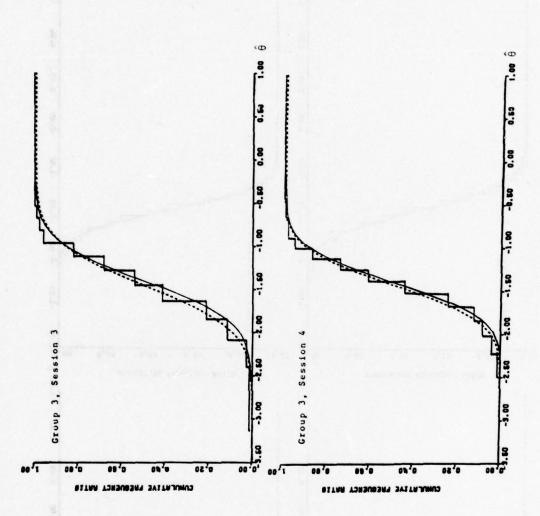


FIGURE A-1-1

Cumulative Frequency Ratio of the Maximum Likelihood Estimates  $\hat{\theta}$  of the 100 Hypothetical Examinees (Solid Line), with the Asymptotic Normal Distribution N( $\theta$ , I( $\theta$ )<sup>-1</sup>) (Solid Curve), and the Normal Distribution Function with the Sample Mean and Variance As Its Two Parameters (Dotted Curve), After Completing Each of the Twenty Sessions:

9 = -1.4

(K. 18)



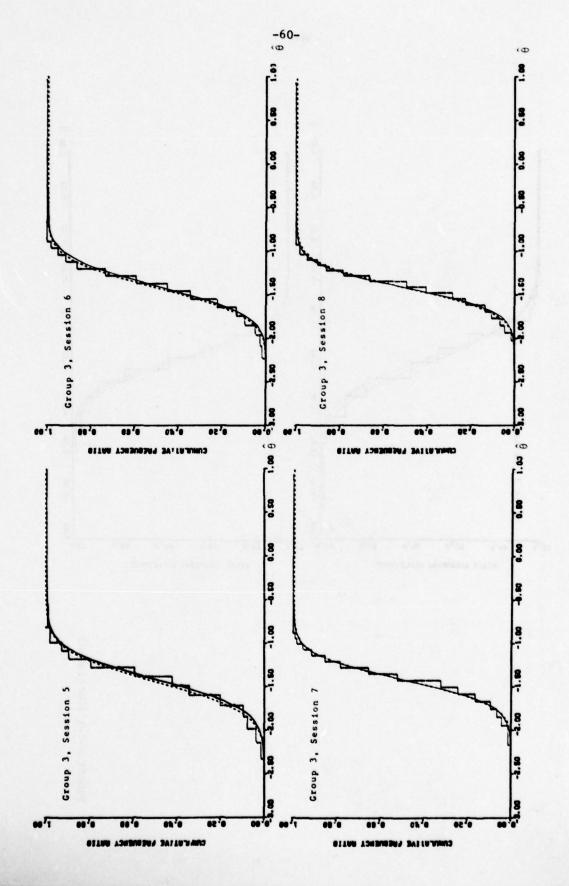


FIGURE A-1-1 (Continued)

FIGURE A-1-1 (Continued)

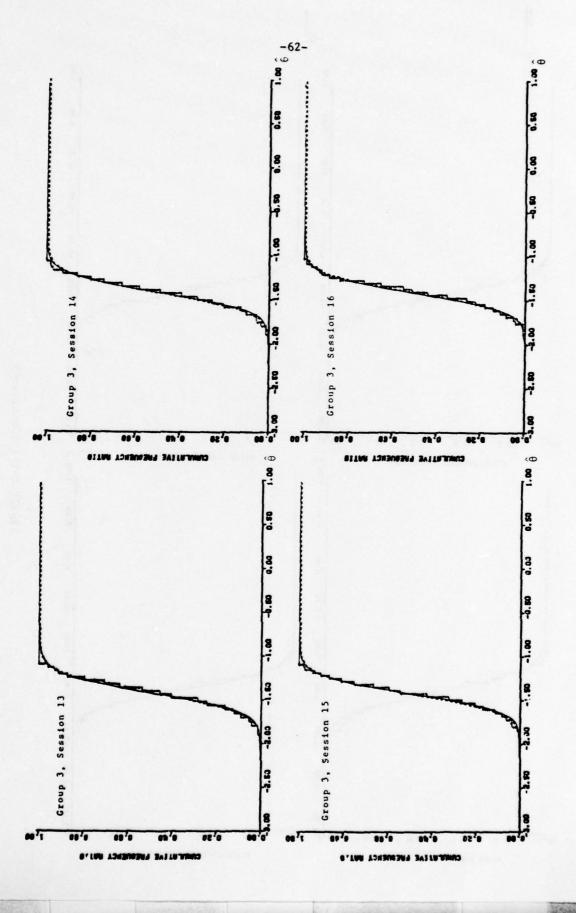


FIGURE A-1-1 (Continued)

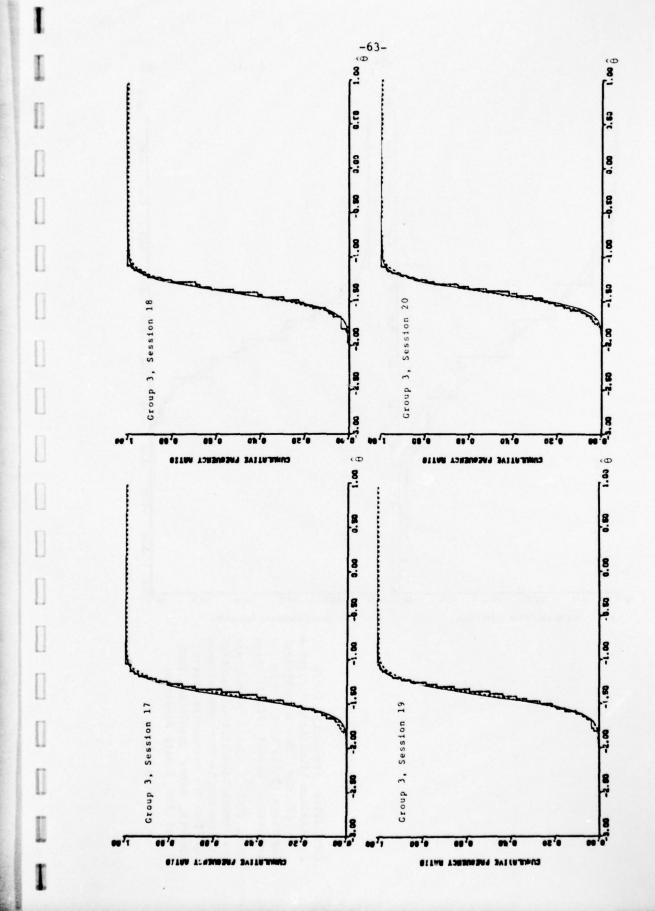


FIGURE A-1-1 (Continued)

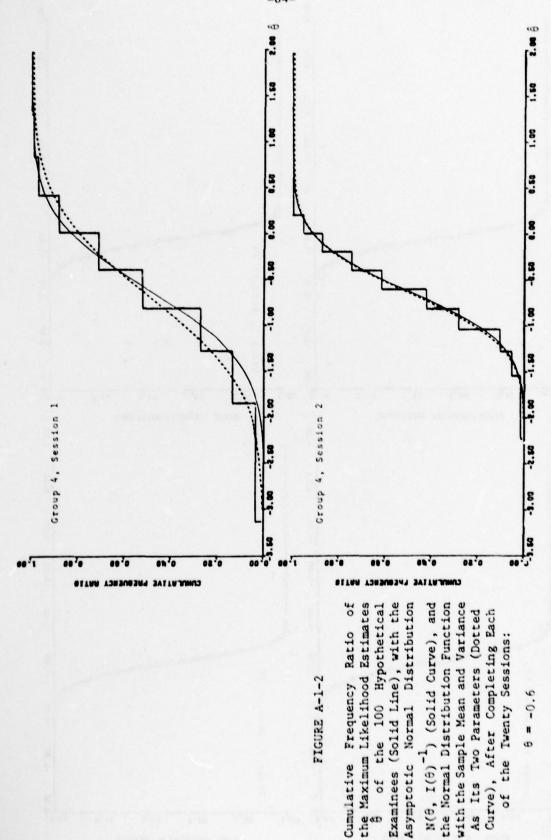


FIGURE A-1-2

Cumulative Frequency

Examinees (Solid Line), with the Asymptotic Normal Distribution  $N(\theta, I(\theta)^{-1})$  (Solid Curve), and the Normal Distribution Function with the Sample Mean and Variance As Its Two Parameters (Dotted Curve), After Completing Each of the Twenty Sessions:

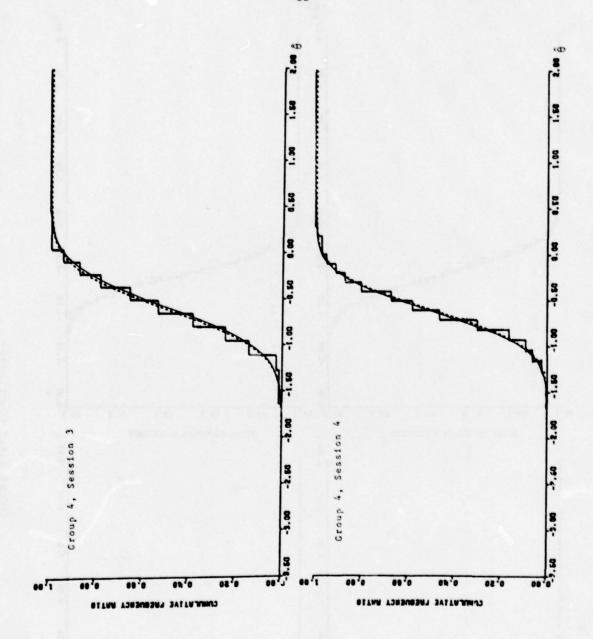
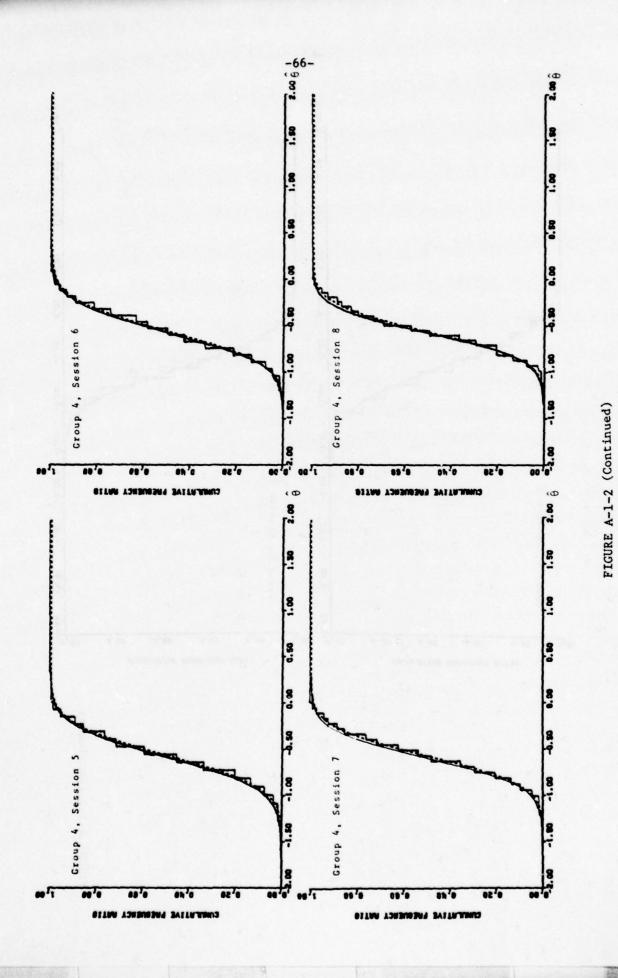


FIGURE A-1-2 (Continued)



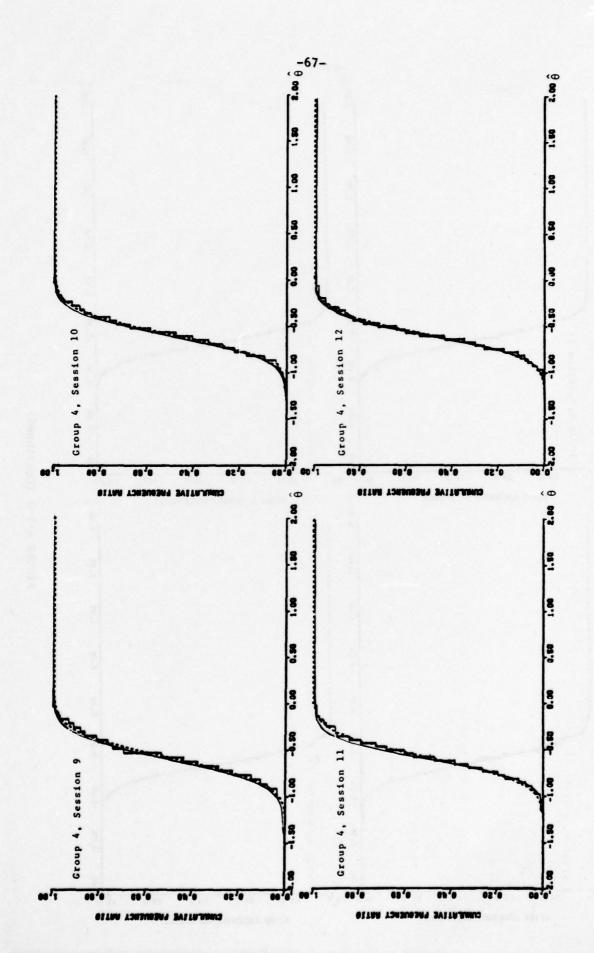


FIGURE A-1-2 (Continued)

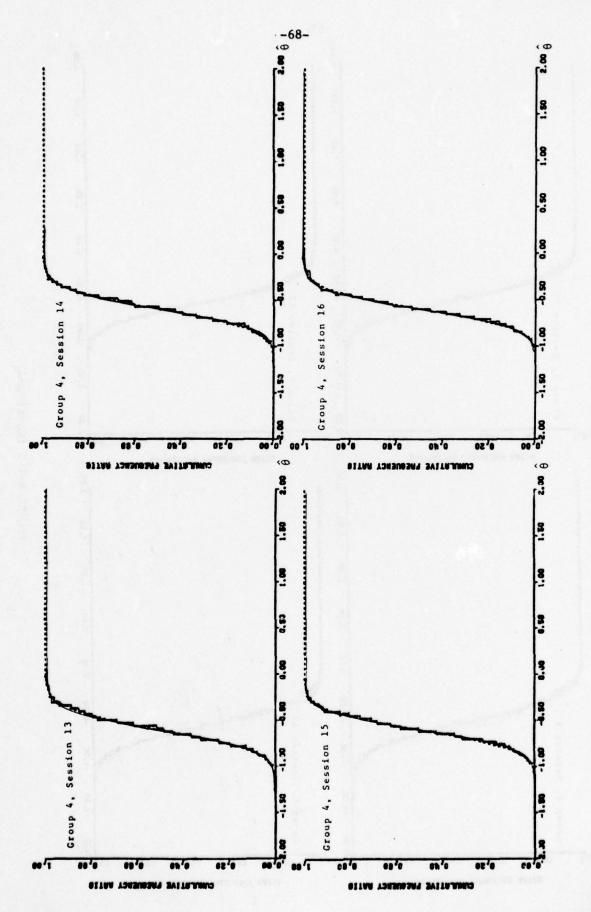


FIGURE A-1-2 (Continued)

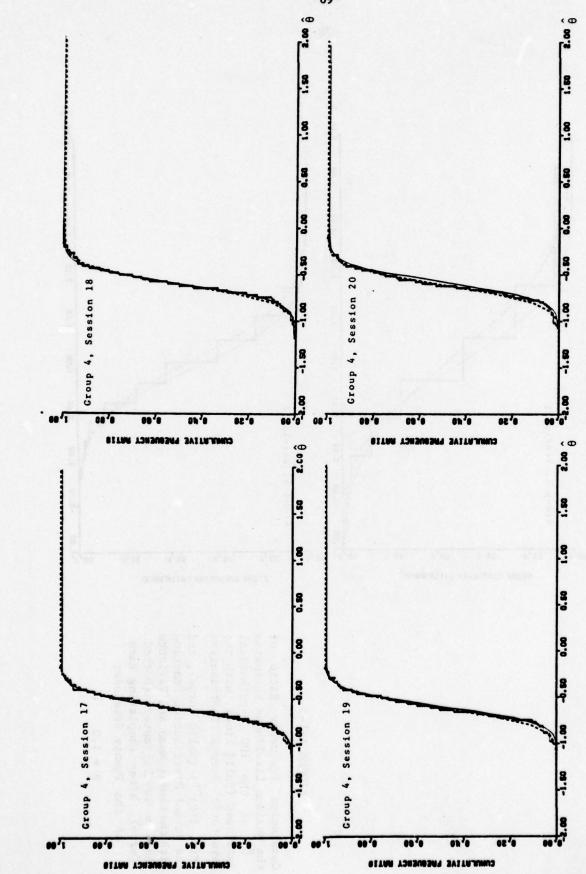


FIGURE A-1-2 (Continued)

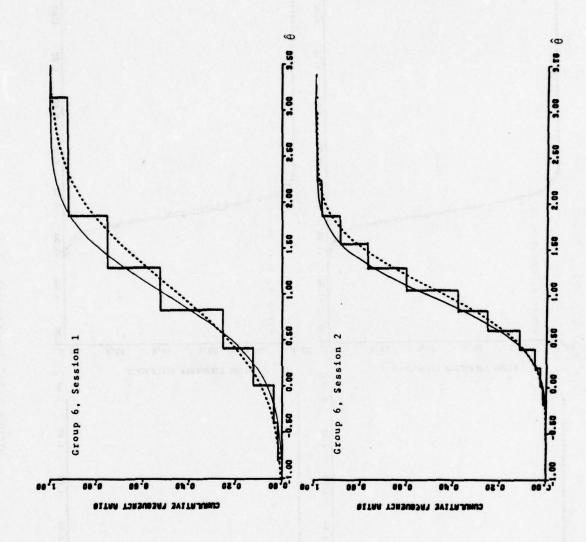
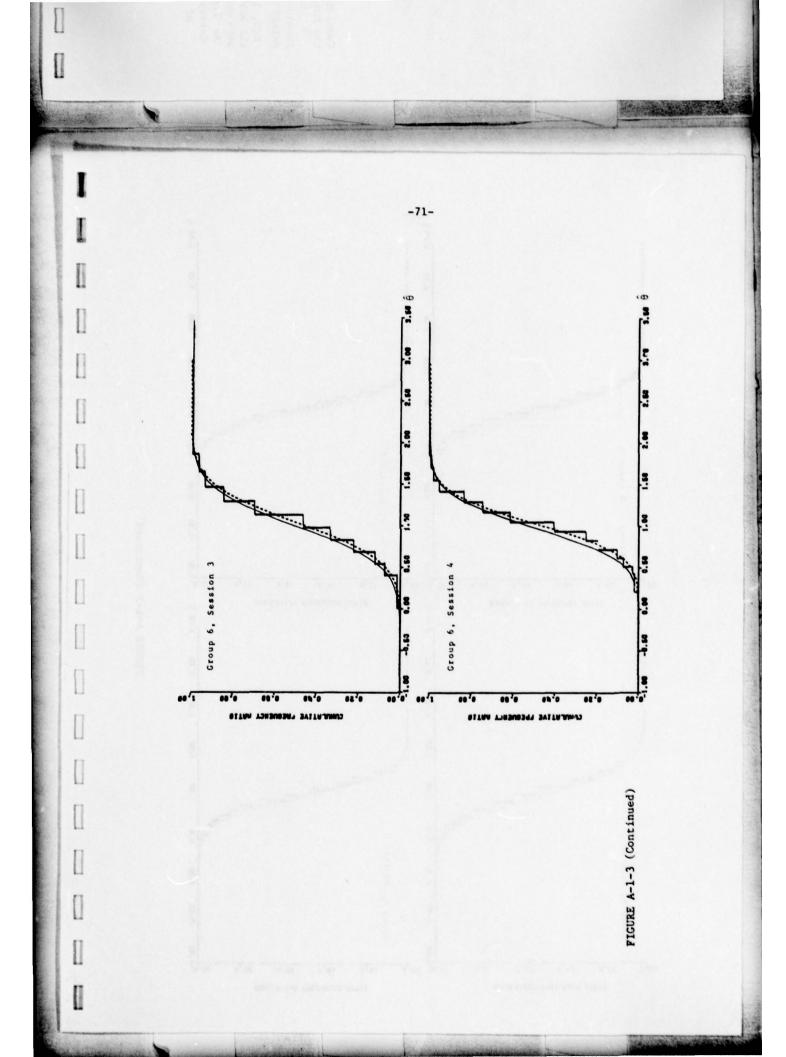
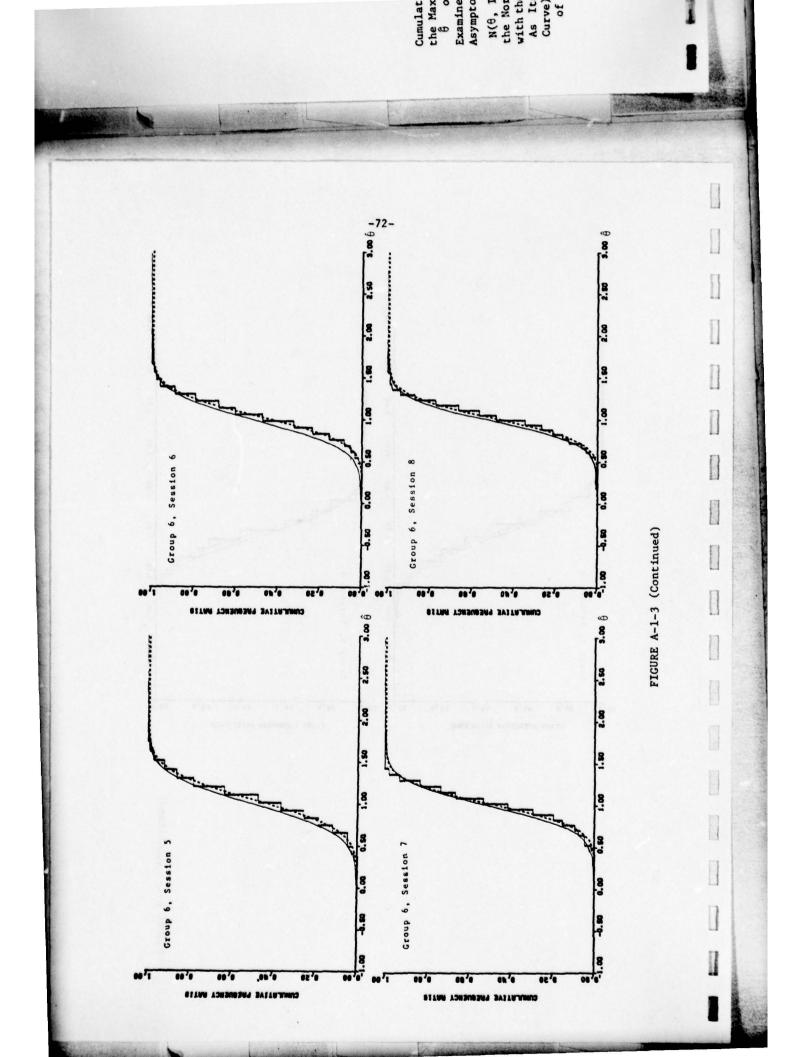


FIGURE A-1-3

Cumulative Frequency Ratio of the Maximum Likelihood Estimates  $\theta$  of the 100 Hypothetical Examinees (Solid Line), with the Asymptotic Normal Distribution N( $\theta$ , I( $\theta$ )<sup>-1</sup>) (Solid Curve), and the Normal Distribution Function with the Sample Mean and Variance As Its Two Parameters (Dotted Curve), After Completing Each of the Twenty Sessions:

3 = 1.0







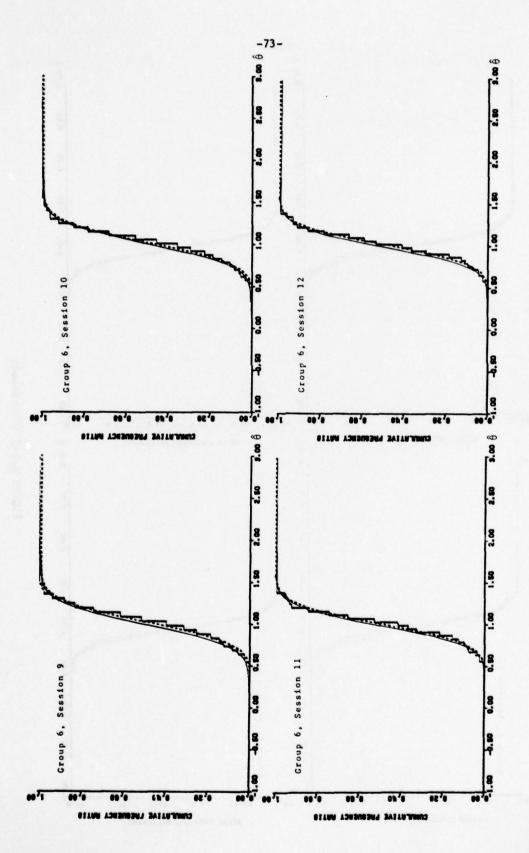
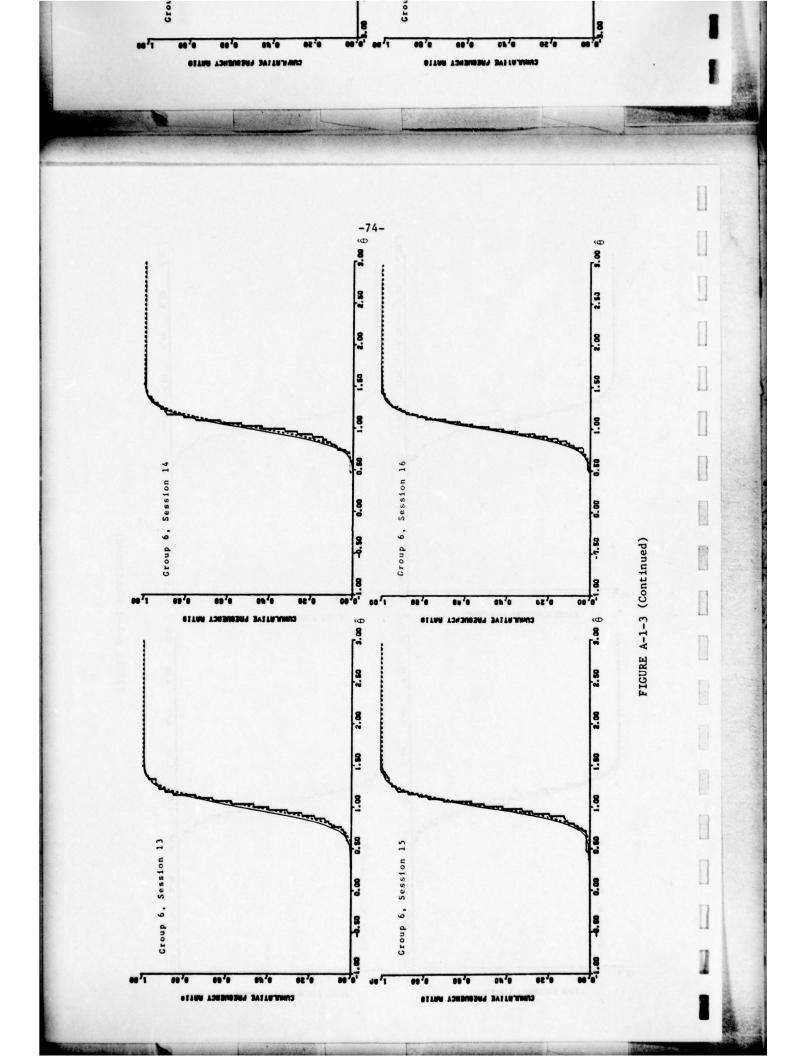
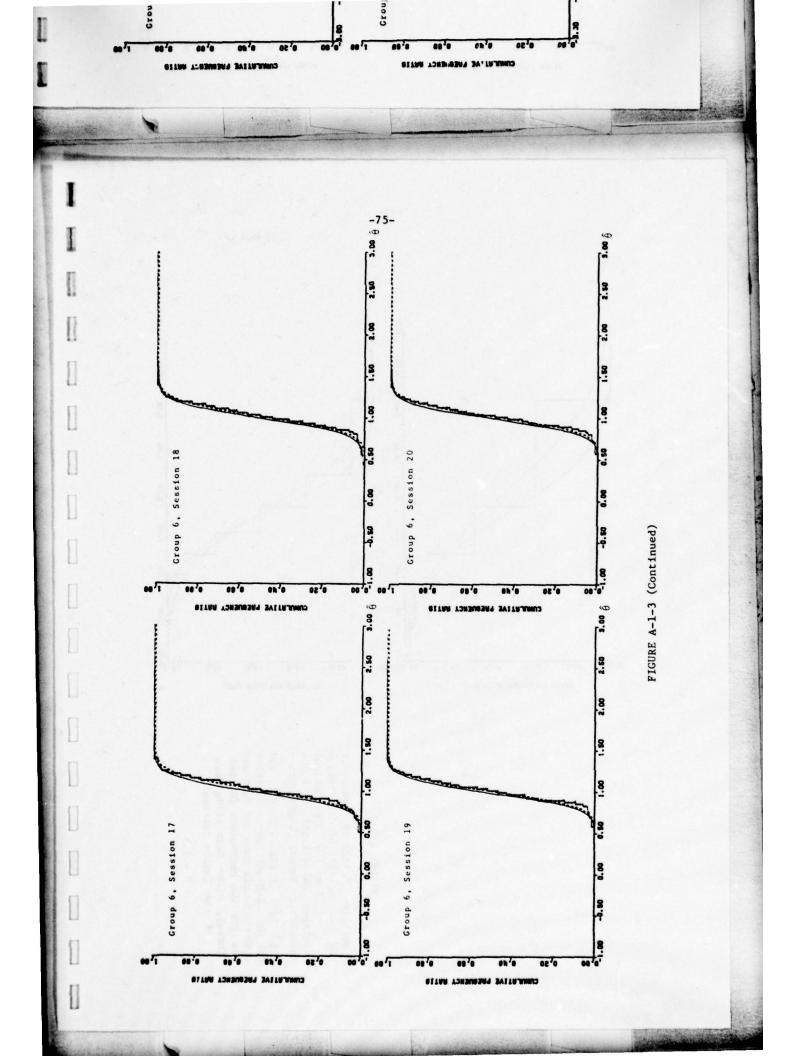


FIGURE A-1-3 (Continued)





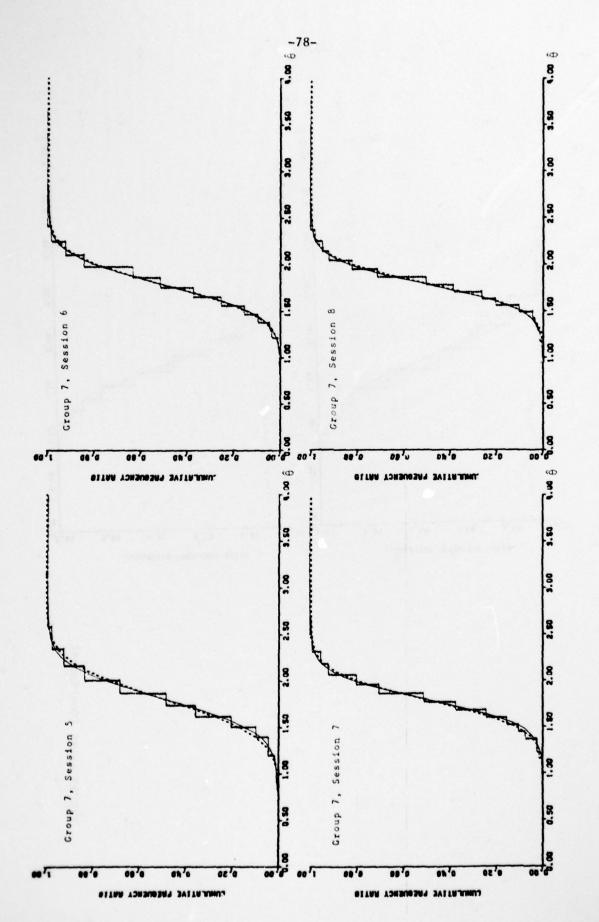


FIGURE A-1-4 (Continued)

FIGURE A-1-4 (Continued)

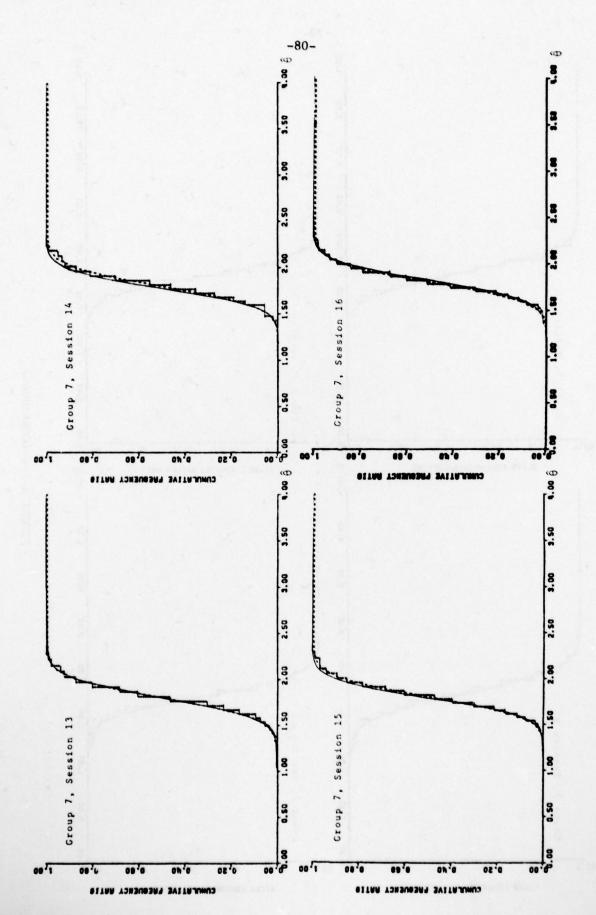


FIGURE A-1-4 (Continued)

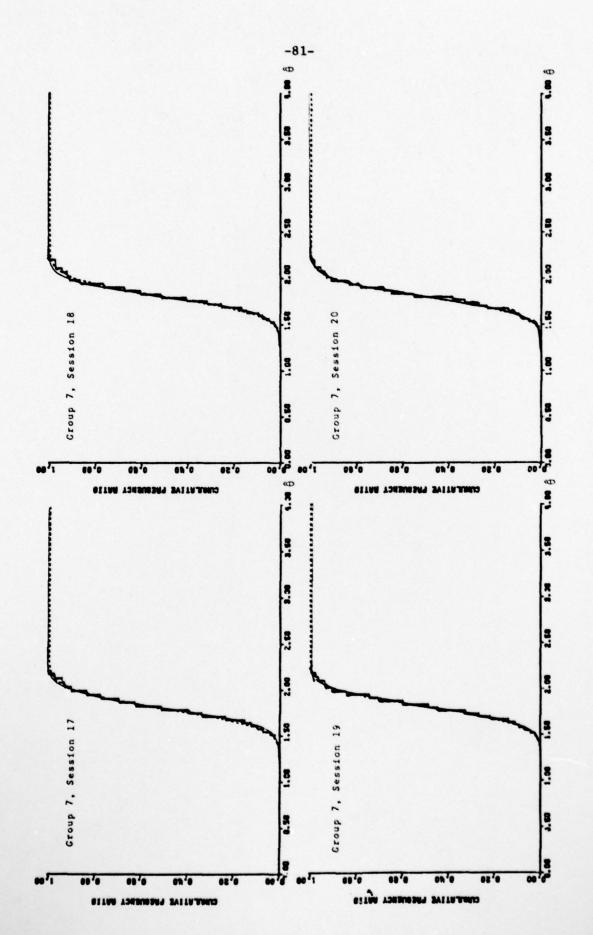


FIGURE A-1-4 (Continued)

APPENDIX II

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TABLE A-2-1

Ten Sample Moments about the Origin of the Maximum Likelihood Estimates of the 100 Hypothetical Examinees after the Completion of Each of the Twenty Sessions. Group 1,  $\theta$  = -3.0.

01	0.6376494, 05 0.6376494, 05 0.6376494, 05 0.6376494, 05 0.6276564, 05 0.657664, 05 0.657664, 05 0.657664, 05 0.657664, 05 0.657664, 05 0.657664, 05 0.65764, 05 0.657644, 05 0.65764, 05 0
٥	0.000 to 100 to
ω	0.94493700 04 0.54493700 04 0.54493700 04 0.94493700 04 0.941588410 04 0.94158410 04 0.94158410 04 0.945600 04
1	10. 20. 20. 20. 20. 20. 20. 20. 20. 20. 2
Moments about Origin 5 6	0.5011/10 03 0.9213/40 03 0.9213/40 03 0.923/2440 03 0.923/2440 03 0.923/2440 03 0.923/2440 03 0.923/2440 03 0.923/2440 03 0.923/2440 03 0.923/2440 03 0.923/2440 03 0.923/2400 03 0.923
Moments ab	-0.306.218.10 03 -0.306.218.10 03 -0.306.218.10 03 -0.306.218.10 03 -0.297.12.20 03 -0.279.14.70 03 -0.279.14.
4	0.57459620 02 0.57479620 02 0.57459620 02 0.57459620 02 0.5954600 02 0.5465490 02 0.5465490 02 0.5465490 02 0.54658100 02 0.54658100 02 0.5468100 02 0.5669550 02 0.5669550 02 0.5669420 02 0.5669420 02 0.5669420 02 0.5669420 02 0.5669420 02 0.5669430 02
3	-0. 310 18340 0.2 -0. 310 18340 0.2 -0. 310 18340 0.2 -0. 310 18340 0.2 -0. 30 37 10 180 0.2 -0. 30 37 25 0.0 -0. 30 25 34 80 0.2 -0. 30 25 34 80 0.2 -0. 20 25 35 10 0.2 -0. 20 25 35 10 0.2 -0. 20 25 35 10 0.2 -0. 20 25 35 10 0.2 -0. 20 25 35 10 0.2 -0. 20 25 35 10 0.2 -0. 20 25 35 10 0.2 -0. 20 25 35 10 0.2 -0. 20 20 0.2 -0. 20 25 35 10 0.2 -0
2	0.94721640 01 0.94721640 01 0.94721640 01 0.94721640 01 0.94721640 01 0.9721640 0
1	0.3142000 01 0.3142000 01 0.3142000 01 0.3142000 01 0.3142000 01 0.314300 01 0.3144700 01 0.314700 01
Session	~~~***********************************

Table A-2-1 (Continued): Group 2,  $\theta$  = -2.2

9 10	-0.18576910 05 0.58273320 05 0.108479320 05 0.21495320 05 0.21495320 05 0.21495320 05 0.21495320 05 0.21495320 05 0.21495320 05 0.21495320 05 0.21495320 05 0.21495320 05 0.21495320 05 0.21495320 05 0.21495320 05 0.21495320 05 0.21495320 05 0.21495320 05 0.21495320 05 0.21495220 05 0.2149520 05 05 0.2149520 05 05 0.2149520 05 05 0.2149520 05 05 0.2149520 05 05 05 05 05 05 05 05 05
æ	0.592 69280 04 0.384 64970 04 0.239 69650 04 0.115 641870 04 0.115 6470 03 0.715 5470 03 0.65 25 5420 03 0.65 25 28740 03 0.65 28740 03
,	-0.116999710 -0.11660160 -0.65341460 -0.65341460 -0.4665520 -0.4665520 -0.4665520 -0.334213630 -0.334213630 -0.3342600 -0.2362600 -0.2362600 -0.2362600 -0.2362600 -0.2362600 -0.2362600
Moments about Origin 5	0.20634250 03 0.2063720 03 0.2063720 03 0.2063650 03 0.10680320 03 0.11462430 03 0.11461730 03 0.11461730 03 0.124740 03 0.124740 03 0.127560 03 0.127560 03 0.127560 03 0.127560 03 0.127560 03 0.127560 03 0.127560 03
Noments ab	10.15.4423320 1.3.4423320 1.3.4423320 1.3.4423320 1.3.4423320 1.3.442320 1.3.
•	20 0 0 0 0 0 0 0 0 0 0 0 0 0 0 0 0 0 0
	10.11.1.1.1.1.1.1.1.1.1.1.1.1.1.1.1.1.1
2	2. 1. 1. 1. 1. 1. 1. 1. 1. 1. 1. 1. 1. 1.
-	25.13.00.00 25.13.00.00 25.13.00.00 25.13.00.00 25.13.00.00 25.13.00.00 25.13.00 25.10.00 25.10.00 25.10.00 25.10.00 25.10.00 25.10.00 25.10.00
Session	-77.00-072-14141414141414141414141414141414141414



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Table A-2-1 (Continued): Group 3,  $\theta = -1.4$ 

	10 0-11314030 05 0-13554730 05 0-1355470 05 0-13
	9 9 9 9 9 9 9 9 9 9 9 9 9 9 9 9 9 9 9
α.	0.12715220 04 0.12716220 04 0.12416250 05 0.7100180 02 0.7100180 02 0.546310 02 0.3175540 02 0.3175540 02 0.3175540 02 0.2677550 02 0.267750 02 0.267750 02 0.267750 02 0.267750 02 0.267750 02 0.267750 02 0.267750 02 0.267750 02
,	10.111.00000 03 10.111.00000 03 10.111.00000 03 10.111.000000 03 10.111.00000 03 10.111.00000 03 10.111.0000 03 10.111.
Moments about Origin 5 6	11 12 2 2 2 2 2 2 2 2 2 2 2 2 2 2 2 2 2
Moments ab	0, 463 44390 02 0, 152 46330 02 0, 102 49990 02 0, 102 49990 02 0, 102 49990 02 0, 102 49990 02 0, 102 418 640 0, 721 6410 01 0, 721 6410 01 0, 721 6410 01 0, 721 6420 01 0, 721 6420 01 0, 668 668 02 0, 668 688 03 0, 668 03 0,
4	0.16483150 02 0.74990803 01 0.5035250 01 0.5655650 01 0.565690 01 0.6570 01 0.465970 01 0.465970 01 0.465970 01 0.469830 01 0.4189360 01 0.4189360 01 0.4189360 01 0.4189360 01 0.4189360 01
3	-0.643 4080 01 -0.397 41130 01 -0.397 41130 01 -0.397 41130 01 -0.397 4120 01 -0.397 420 01 -0.397 420 01 -0.297 430 01 -0.297 430 01 -0.297 7810 01
2	0.22777230 0.2177230 0.2177230 0.2177230 0.2177230 0.2177230 0.2175250 0.217
1	0.14534910 01 0.14508710 01 0.14508710 01 0.14508710 01 0.14158400 01 0.14158400 01 0.14158400 01 0.14158400 01 0.14158800 01 0.14158800 01 0.1416800 01 0.1416800 01 0.1416800 01 0.1416800 01 0.1416800 01 0.1416800 01 0.1416800 01 0.1416800 01 0.1450800 01 0.1450800 01 0.1450800 01
Session	

Table A-2-1 (Continued): Group 4,  $\theta = -0.6$ 

Session					Moments about Origin	ut Origin					
	-	2	F	7	5	9	7	80	6	10	
	0.67866000 00 0.65315000 00 0.65315000 00 0.554200 00 0.554200 00 0.554200 00 0.554200 00 0.554200 00 0.554300 00 0.554300 00 0.5555000 00 0.5555000 00 0.5555000 00 0.5555000 00 0.5557000 00	0.111776 u 0.5431470 0.5436430 0.3551270 0.3551270 0.3551270 0.3551270 0.3165320	-0.19807440 01 -0.1047170 00 -0.4821950 00 -0.3029190 00 -0.3028190 00 -0.2028180 00	0. 4650642D UI 0. 53079940 00 0. 526 264D 00 0. 259 464D 00 0. 259 464D 00 0. 259 765D 00 0. 205 765D 00 0. 196 784D 00 0. 196 784D 00 0. 196 784D 00 0. 198 789D 00 0. 188 789D 00 0. 189 108D 00	0.11929910 U2 0.13961690 U1 0.34262720 00 0.34262710 00 0.2255760 00 0.22557760 00 0.129718920 00 0.159718920 00 0.159718920 00 0.159718920 00 0.159718920 00 0.159718920 00 0.159718920 00 0.159718920 00 0.159718920 00 0.159718920 00 0.159718920 00 0.159718920 00 0.159718920 00 0.159718920 00 0.159718920 00 0.159718920 00 0.159718920 00 0.159718920 00	0.255910.10 0.255910.10 0.255910.10 0.255910.10 0.1250.10 0.1150.20 0.1160.20 0.1160.20 0.1260.0	0.4573590 00 -0.4642090 00 -0.464600 00 -0.464600 00	0.85474550 03 0.85474550 03 0.10561520 01 0.2163440 00 0.1163440 00 0.116340 00 0.1266070 00 0.1266070 00 0.1266070 00 0.1266070 00 0.12650070 00 0.12650070 00 0.1263400070 00 0.1263400070 00 0.1263400070 00 0.1263400070 00 0.1263400000 00 0.126340000 00 0.126340000 00 0.126340000 00 0.126340000 00 0.12634000 00 0.1263400 00 0.1263400 00 0.1263400 00 0.126340 00 00 00 00 00 00 00 00 00 00 00 00 00	- 6,422 625 0 0 0 0 0 0 0 0 0 0 0 0 0 0 0 0 0 0 0	0.200.255.04 0.207.255.04 0.207.200.02 0.207.200.00 0.172.456.00 0.172.456.00 0.172.456.00 0.172.456.00 0.007.200.00 0.007.200.00 0.007.200.00 0.007.200.00 0.007.200.00 0.007.200.00 0.007.200.00 0.007.200.00 0.007.200.00	



PROM CONY PURE BEST QUALITY PRACPICATES

Table A-2-1 (Continued): Group 5,  $\theta = 0.2$ 

	10	10 0210120 01	0.43753413-01	0.10194910-01	0.10657136-02	0.4515465-03	0.11901790-03 C. e65504c0-04 O.042c43yp-04	0.43113650-04
	o	0.12556520 01	0.79396395-31	0.10.4911.0	0.25504420-02	0.12505273-02 0.14173723-02 0.5005000-02	0.1725000-03 0.1725000-03 0.13373030-03	
·	•	0.14061620 00	0.31037250-01	0.23144740-02	0.331136.00-02	0.7520000-03	0.36293270-03 0.26388910-03 0.20768500-03	0.19703170-03
,		00.014524000	0-334333/0-01	0.40102110-02 0.40102110-02	20-040400040-0	0-1400+50-02 0-10041710-02 0-93912520-03	0.1017/0300103	0
it Origin 6		0.15/60/60 00	0.23464210-01 0.22161200-01 0.12634741-01	0.67011640-02	0.4165.050-02 0.47934320-02	0.19525015-02 0.19525015-02	0.1386caub-02 0.1144co4p-02 0.110949cb-02	0.107213-0-02
Moments about Origin 5	0.512*3670 00	0.13039620 00 0.0c161450-01 0.45612010-01	0.29051360-01	0-1041490-01	0.1202.746J-01 0.425517J-02 0.5496206J-02	0.44037020-02	0.32370355-02 0.2522050J-02 0.27408740-02 0.2839145J-02	0.26552255-02
4	0. 00184050.00	0-10252c10 00 C-02703710-01	0.41544565-01	0.1575559-01	0.15251cau-01 0.11471060-01	0.0225920-02	0.1270010-02 0.1120400-02 0.13472540-02	70-09074769-0
3	0. 26130560 00	0.10685610 00 0.774961000	0.5373848C-01 0.4165746U-01	0.36127420-01 0.36127420-01 0.35436430-01	0.25633710-01	0.221945456-01	9.19e5476U-01 0.19e25275-01 0.2019564U-01	
8	0. 0.01516.00	3-14779333 00 0-14033030	0-11137660 00 0-3-2-1700-01	0.019/23/0-01	0.0000000000000000000000000000000000000	0.0000400-01	0.09459446001	
1	3.413//606 03		0. 40494 030 00	0.15 14 1030 00 0.15 14 1030 00	3-1463-000 00 0-16771030 00	0.19564000 00	0.2050c000 0.2050c000 0.2050c000	
Session		•••	- 0	*9:	222	247	110	

- T. 19.

Table A-2-1 (Continued): Group 6,  $\theta = 1.0$ 

Session					Moments about Origin	it Origin				
	1	7	3	4	\$	9	,	80	6	10
-	0.11410300 61	3.1.557.00 01	0.41090310 01	0.10573110 62	0.29142865 02	d. a. 5030320 62	4.6.3.4.00	A 1000 1000 1000 1000 1000 1000 1000 10		
7	0.10023330 01	0.1.2000-1.0	0.22169980 01	10 00161676.0	0.7562510.01	J. 18190160 02	60 60 60 60 60	000000000000000000000000000000000000000	חפונוניים	6. (200445) 04
•	0-104000000	3.14.74333 91	0.15678250 01	0.41498140 31	0.31372340 01	20 000000000000000000000000000000000000	30 5553555		0.2250/// 03	0.105.0523 04
•	3.1342.703 01	3.117.13.10 01	0.14054650 01	0.17674170 01	0.2421277.01	10 000000000000000000000000000000000000			0.25235310 32	0.41551080 02
•	3-13-5-530 01		13701470	3-16974. 6.1 01		10 010000	10 07/05944-0		0.5444.60 01	0.15212553 02
•	0.10272-00 01	J	0.14940100 01	0.1522.00	10 09405 311 0	10 055 70187.0			0.70504000 01	Collegeled Oc
1	0-10455550 01			0-156-7110 01		0.15366130	0.30143000 01		0.55657070 01	0.73742920 01
			0.12115930 01	0.1.27.57.00	10 001011010	חיווינים וויים			0.30145100 01	0.45725000 01
*	0.10254400 01			0.1045540	0 1500000		10 (240)0200	0.265-7930 01	6.36452300 01	0.40625520 01
07			1855350	0 0315255150		10 00000000	3.21200213 01		0.321025cu 01	0.39931900 01
11			0.11572620 01	10 01604010	10 000000000	0.1811.700 01	10 000000177-0	0.20081320 01	0.3.353510 01	0.42405763 01
77	0.10191000		0-11553500 01	10 0100000000		0.1/1/25470 01	0.10.010.01.0		0.21393703 01	0.40207930 01
13			0-11-34160 01	10 00001710	10 001800010		0.1535001. 01		0.300000000	0.30003300 01
1.			0.11321450 01	0.15.357570		0.15736770 01			0.20902220 01	0.33361550 01
15	6. 10154430 01		0-11214560 01	0.121631561			10 04040601.0		0.20732362 01	0.33316110 01
97	0.10127200 01	0.1050-040	10 07 57 5 1 1 1 0	0.1204740	10 0511750		10 00000000000		0.24433020 01	0.2569952.0
11	3.10160230 61		0.111654111.01	0.12029:10			0.11166363 01		0.23823.12 01	10 01824702.0
10	0.10100100	3.103.61.00	0.11145700 01	10 0106 707110			10 0-1/0001-0		0.21505550 01	0.25746020 01
61	0.10198700 01	13	0-11227500	10 0011001110	0.13033610 01	10 00+0764100	10 01001-01-0	0.15621730 01	0.21464510 01	0.25075550 01
07	d-10201400 01	3.10.4444.01	11160130 01	10 0001100	10 0553535100	0.14/30430 01	10 0000001.0	0.14505590 01	0.21544410	10 0354565.0
			10 000000	0.12010020 01	0.13031130 01	0.144/4090 01	J.10234752 01	0.10454000 01	0.21233500 01	0.24719763 01



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Table A-2-1 (Continued): Group 7,  $\theta = 1.8$ 

		36444383838383383833 49393934333333338833
	10	0.000   0.000
		111.4.1.1.2.4.1.1.2.4.1.1.2.4.1.1.2.4.1.2.
	•	0.1012241 0.101224 0.1072724 0.1072727 0.1072727 0.1072777 0.10727777 0.107277777 0.10727777777777777777777777777777777777
		0.12581310 04 0.1225760 04 0.41841210 05 0.41841210 05 0.41841210 05 0.14841210 05 0.14841210 05 0.14841210 05 0.1496413 0
	80	
	1	111420000000000000000000000000000000000
ıl.	9	13.500.2556 13.50
out Orig		11 1 1 1 1 1 1 1 1 1 1 1 1 1 1 1 1 1 1
Moments about Origin	•	0.12069390 03 0.8364649 02 0.8364669 02 0.836460 02 0.82666510 02 0.82666510 02 0.82666510 02 0.82666510 02 0.82666510 02 0.82667510 02 0.82667510 02 0.8267510 02
Mor		0.12009350 0.55464850 0.53464850 0.53454500 0.22656510 0.22640550 0.22620010 0.21620010 0.206210 0.206210 0.206210 0.206210 0.206310 0.2063210
		20000000000000000000000000000000000000
	4	0.000 444 400 600 600 600 600 600 600 600
		25,000 00 25,000 01 27,000
	3	0.012970 0.01297 0.012
		553535353535353535355 2222323232323233323
	2	10.00000000000000000000000000000000000
		555555555555555555555555555555555555555
	1	0.1801.4500 0.1811.4500 0.1811.4500 0.1811.4700 0.1804.600 0.1774.7700 0.1774.7700 0.1774.7700 0.1774.7700 0.1801.600 0.1
Conston	norses of	

	-91-
:	0.7044307 05 0.6235313 05 0.51442130 05 0.5235022 05 0.1337020 05 0.1322450 05 0.1222450 05 0.2222450 05 0.225260 05 0.2559640 05 0.2312060 05 0.2312060 05 0.172240 05 0.172240 05
٠	0.224+5440 05 0.20121730 05 0.16641330 05 0.16641330 05 0.154420 05 0.1037440 05 0.5374410 05 0.5374410 05 0.5574410 05 0.
w	0.7157550 04 0.64597550 04 0.6488300 04 0.6488720 04 0.6488720 04 0.376420 04 0.37690 04 0.277270 04 0.277280 04 0.277280 04 0.277280 04 0.277280 04 0.2872820 04
,	0.22854740 04 0.28854740 04 0.1595550 04 0.1595500 04 0.1587710 04
ut Origin 6	0.73142700 03 0.5834650 03 0.58445310 03 0.51996710 03 0.475445310 03 0.4056450 03 0.4056450 03 0.4194670 03 0.41184670 03 0.41184670 03 0.451760 03 0.3517670 03 0.3517670 03 0.3517670 03 0.3517670 03 0.3517670 03 0.3517670 03 0.3517670 03 0.3517670 03 0.3517670 03
Moments about Origin 5 6	0.234,704,0 0.21495,504,0 0.21495,504,0 0.30,1719,504,0 0.30,1719,504,0 0.30,1719,504,0 0.30,1717,717,0 0.30,1749,504,0 0.30,1
7	6. 724 2410 0 0 0 0 0 0 0 0 0 0 0 0 0 0 0 0 0 0
	0.248170;0 02 0.2405090 02 0.21518650 02 0.41518650 02 0.4157090 02 0.4157090 02 0.41590460 02 0.41590460 02 0.41590460 02 0.41590460 02 0.41590460 02 0.4159130 02
2	10 0.124.2020 10 0.1
	0. 241-27 D. 0. 24
Session	

 $\theta = 2.6$ 

Table A-2-1 (Continued): Group 8,

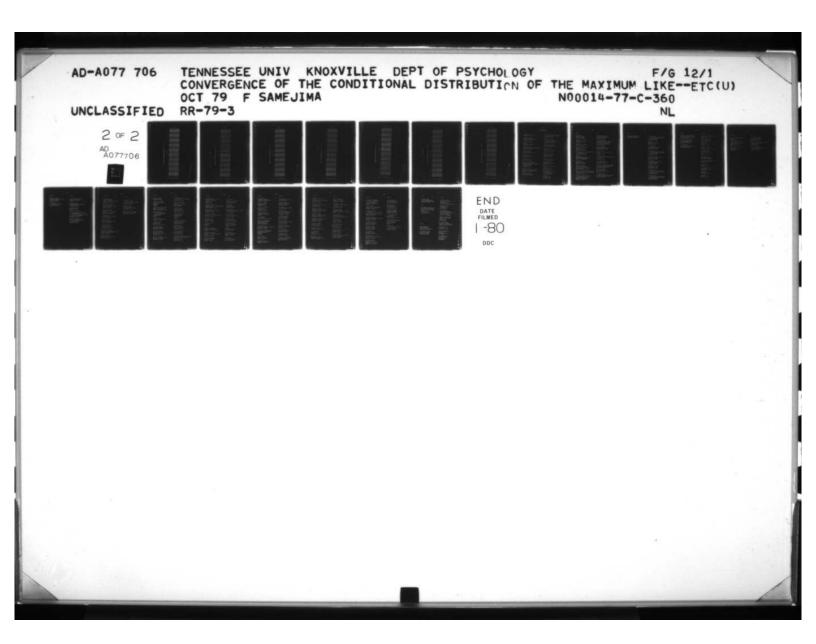
PROM COLY FURBISHIND NO LING

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TABLE A-2-2

Ten Sample Moments about the Mean of the Maximum Likelihood Estimates of the 100 Hypothetical Examinees after the Completion of Each of the Twenty Sessions. Group 1,  $\theta$  = -3.0.

10	0.0 0.0 0.0 0.0 0.0 0.0 0.0 0.0 0.0 0.0
ø	0.00 0.00
<b>x</b> 0	0.0 0.0 0.0 0.0 0.1759.6470-0 0.1759.74-0 0.1759.74-0 0.1759.74-0 0.1759.74-0 0.1759.75-0 0.1759.0 0.1759.0 0.1759.0 0.1759.0 0.1759.0 0.1759.0 0.1759.0 0.1767.0 0.1
,	0.00 0.00 0.00 0.00 0.00 0.00 0.00 0.0
t Mean 6	0.00 0.00 0.00 0.00 0.00 0.00 0.00 0.0
Moments about Mean 5	0.0 0.0 0.0 0.0 0.0 0.0 0.0 0.0 0.0 0.0
7	0.0 0.0 0.0 0.0 0.0 0.0 0.0 0.0 0.0 0.0
9	0.0 0.0 0.17778960-02 0.17778960-02 0.59478020-02 0.59480020-02 0.37880020-02 0.37880020-02 0.37880020-02 0.37880020-02 0.3813140-02 0.38185620-02 0.3841310-02 0.3841310-02 0.3841300-02 0.3841300-02
2	0.0 0.0 0.0 0.19357au-02 0.119357au-02 0.129345au-01 0.129345au-01 0.129345au-01 0.1352010-01 0.1352010-01 0.1352010-01 0.15522au-01 0.15522au-01 0.119479u-01 0.119479u-01
-	0.0 0.0 0.0 0.0 0.0 0.0 0.0 0.0 0.0 0.0
Session	-70,70,00000000000000000000000000000000



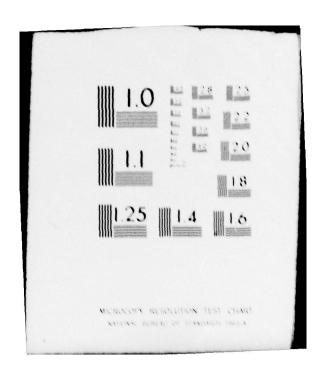


Table A-2-2 (Continued): Group 2,  $\theta = -2.2$ 

To and the last

01	0.01 0.25981900 0.2 0.25981900 0.2 0.2 0.2598190 0.1 0.1 0.2 0.2 0.2 0.2 0.2 0.2 0.2 0.2 0.2 0.2
۰	0.1228/33U 02.0.117462D-01.0.2117462D-01.0.2117462D-01.0.2117462D-01.0.2117462D-01.0.2117462D-01.0.2117462D-01.0.2117462D-01.0.2117462D-01.0.2117462D-01.0.2117462D-01.0.2117462D-01.0.211752D-01.0.2117
ω	0. 5946.2380 01 0. 812.93720-01 0. 11.18950-01 0. 15.1820.050-01 0. 15.1820.00 0. 205.7570-03 0. 209.8900-03 0. 299.8900-04 0. 298.8460-04 0. 377.87480-04 0. 598.8460-04 0. 598.8460-04 0. 298.8460-04 0. 298.8460-04 0. 298.8460-04 0. 298.8460-04 0. 298.8460-04 0. 298.8460-04 0. 298.8460-04 0. 298.8460-04 0. 298.8460-04
-	0.299646.0 0.3206646.0 0.3208646.0 0.310318684.0 0.1729490.0 0.420345.0 0.420345.0 0.22771410.0 0.22771410.0 0.237714100.0 0.237714100.0 0.237714100.0 0.237714100.0 0.237714100.0 0.237714100.0 0.237714100.0 0.237714100.0 0.237714100.0 0
out Mean 6	0.15823465 01 0.10598840 00 0.10598840 00 0.55638000-01 0.112822800-01 0.112822800-01 0.11282800-02 0.128810-02 0.12810-03 0.28158610-03 0.28158610-03 0.21858610-03 0.21858610-03
Moments about Mean 5	0.05845050 0.1014484-01 0.1014184-01 0.1014184-01 0.1014184-01 0.1014182-01 0.10
4	0.6x441-370 0.11146110 0.11146110 0.2x2410-01 0.2x27410-01 0.2x272410-01 0.2x272410-01 0.2x272410-01 0.2x272410-01 0.2x272410-02 0.2x272410-02 0.2x272410-02 0.2x27210-02
r	0.530.22159 0.530.22159 0.5446.540.00 0.5446.540.00 0.5466.00 0.5466.00 0.5
7	10-1-1-1-1-1-1-1-1-1-1-1-1-1-1-1-1-1-1-
1	0.139920-1-1-1-1-1-1-1-1-1-1-1-1-1-1-1-1-1-1-1
Sestion	

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Table A-2-2 (Continued): Group 3,  $\theta = -1.4$ 

1	Seseton					Moments about Mean	out Mean				
0.47919990 06 -0.41484140 00 0.1300200 01 0.1722990 06 -0.413790-01 0.1252310 00 0.10879120 06 -0.1447920-01 0.111111910 00 0.10879120 06 -0.1447920-01 0.15171910 00 0.10879120 01 -0.1456220-01 0.1517430-01 0.4127710-01 -0.145620-02 0.1517430-01 0.5514870-01 -0.4178510-02 0.15174310-02 0.46471120-01 -0.4178510-02 0.177310-02 0.46471120-01 -0.4178510-02 0.177310-02 0.46471120-01 -0.4178510-02 0.177310-02 0.46471120-01 -0.4178510-02 0.177310-02 0.27490-01 -0.4178510-02 0.177310-02 0.27490-01 -0.4178510-02 0.1773130-02 0.274700-01 -0.417310-02 0.1773130-02 0.274700-01 -0.417310-02 0.1773130-02 0.274700-01 -0.417310-02 0.1773130-02 0.274700-01 -0.417310-02 0.1773130-02 0.274700-01 -0.417310-02 0.1773130-02 0.274700-01 -0.417310-02 0.177310-02 0.274700-01 -0.417310-02 0.177310-02 0.274700-01 -0.417310-02			7		•	5	e	7	100		10
0.1725390 00 -0.3113702-01 0.2152310 00 0.1725310 00 0.1025390 00 -0.175622-01 0.111111910 00 0.10253320 00 -0.1756222-01 0.11111910 00 0.10253320 00 0.17562220-01 0.15174370-01 0.13174370-01 0.13174370-01 0.15174370-01 0.15174370-01 0.15174370-01 0.15174370-01 0.15174370-01 0.15174370-01 0.15174370-01 0.15174370-02 0.175740-02 0.175740-02 0.175770-01 0.175770-02 0.175770-02 0.175770-02 0.175770-02 0.175770-02 0.175770-02 0.175770-02 0.175770-02 0.175770-02 0.175770-02 0.175770-02 0.175770-02 0.1757770-02 0		-0.14521720-14	0.05535980 00	-0.41486540 00	0.13002090 61	-0.151 235 20 01	A				
0.1722890 00 -0.15447320-01 0.11111910 00 0.10873120 00 0.1187120-01 0.1567220-01 0.1667270-01 0	,	-3.435207.0-15	0.4435550 00	-0-34139900-01	0.21523710 60	20 114 11900 00	G. 52173 520		0.31-16210 01	-0.137 occito	0.254costu uz
0.1037320 00 -0.155 6720-01 0.465329-01 0.4213730-01 0.42137320-01 0.42137320-01 0.42137320-01 0.42137320-01 0.42137320-01 0.42137320-01 0.42137320-01 0.42137320-01 0.42137320-01 0.42137320-01 0.42137320-01 0.42137320-02 0.42137320-02 0.42137320-02 0.42137320-02 0.42137320-02 0.42137320-02 0.42137320-02 0.42137320-02 0.42137320-02 0.42137320-02 0.42137320-02 0.42137320-02 0.421373230-02 0.42137320-02 0.4213720-02 0.4213720-02 0.4213720-02 0.4213720-02 0.4213720-02 0.4213720-02 0.4213720-02 0.4213720-0	•	+0-16565920-0-		10-02666915-0-	0-11141910 00	20 14254340 00	to the first to the	20 0010100000	4.45511620 OF	-C. 11 9 de 11	0.43122840 01
0.4137730-01 -0.1286225-01 0.22906460-01 0.23907480-01 0.23907480-01 0.23907480-01 0.23907480-01 0.23907480-01 0.23907480-01 0.23907480-01 0.23907490-02 0.23907490-02 0.23907490-02 0.23907490-02 0.239077100-02 0.2390		-0.14722750-14		-0.15466721-01	0.46632930-01	-0. 149 11343-01		200 00000000000000000000000000000000000	0.6000000000	-0.113-0370 01	0.19257360 61
0.565348710-01 0.659876850-02 0.15374570-01 0.565348710-02 0.565348710-02 0.15374570-01 0.565348710-02 0.15374570-01 0.565348710-02 0.15474570-02 0.564471320-02 0.5647174710-02 0.5647174710-02 0.5647174710-02 0.5647174710-02 0.564717470-02 0.564717470-02 0.564717470-02 0.576732	•	+1-019704410-14	0.41347930-01	-6-12566259-31	0.22904680-01	-0.9034130.00	J. 147 sh. st. etc.		10-030,000000	-0.146+14301	3.27.01115-01
0.5332710-01 0.5683710-02 0.1049130-02 0.46471120-02 0.46471120-02 0.4175340-02 0.46471120-02 0.4175340-02 0.46471120-02 0.4175340-02 0.46471120-02 0.4175340-02 0.46471120-02 0.4175340-02 0.46471120-02 0.4175340-02 0.46471120-02 0.41754440-02 0.41754400-02 0.4175440-02 0.4175440-02 0.4175440-02 0.4175440-02 0.4175440-02 0.4175440-02 0.4175440-02 0.4175440-02 0.41754400-02 0.41		-0.10763610-14	0.05150250		0.15474670	20 676 - 10 - 01			75-03104160-05	-0-101727-07	610cerio-0c
0.56554870-01 0.51171839-02 0.78779340-02 0.4521770-01 0.457279-02 0.5789860-02 0.45421726-01 0.4552470-02 0.5789860-02 0.45421726-01 0.4552470-02 0.4542860-02 0.4542860-02 0.4542860-02 0.4542860-02 0.4542860-02 0.4542860-02 0.4542860-02 0.4542860-02 0.45428	1	-9-17399370-1-	0.53832710-61		0 1070 1530	79-7967500	20.000000000000000000000000000000000000		0.355,035,05	-0.cic.5.co-02	0.1s03775.0-02
0.45237700-01 -0.41785610-02 0.5184860-02 0.464511120-01 -0.4522410-02 0.5184860-02 0.46451120-01 -0.4522410-02 0.5184860-02 0.46451120-01 -0.4522410-02 0.5184860-02 0.464512		-0-12823030-16	0.50514975-31		201161010	20-03-03-05	25. 200 10 10 10		0-15174870-02	-6. yearing a-1.	3.700+7.20-03
0.45451120-01 0.3452419-02 0.5788869-02 0.351744-01 0.1226429-02 0.4543869-02 0.351744-01 0.11226429-02 0.4543890-02 0.2743220-01 0.11226429-02 0.2547100-02 0.2743220-01 0.11326429-02 0.2547100-02 0.2743220-01 0.1132659-02 0.2547100-02 0.2743230-01 0.1132669-02 0.1123829-02 0.2547400-01 0.1157966002 0.118773970-02 0.2547400-01 0.1157960002 0.11877970-02 0.2547400-01 0.1157960002 0.11877970-02	•	-0-14223340-14	0.45233750-01		20.000000000000000000000000000000000000	20-021906970-02	70-00000000		0-45235460-03	-0.201760-02	0-1+600000000000000000000000000000000000
0. \$60.21\$50-01 -0.31265420.02 0.4250620-02 0.317442-01 -0.27711230-02 0.32597390-02 0.31777100-02 0.317777100-02 0.317777100-02 0.317777100-02 0.317777100-02 0.317777100-02 0.317777100-02 0.31777720-02 0.3177777100-02 0.31777777777777777777777777777777777777	10	-0.16256550-14	0.46451120-01		20-011-1110-0	20-0-806021-0-	20-02/02/01/0		1.21045460-03	-0.10503524-54	0.45021Bc04
0.3614444-01 -0.1722459-02 0.32597390-02 0.47632 20-01 -0.2711429-02 0.22497700-02 0.47632 20-01 -0.1934639-02 0.224977700-02 0.47632 20-01 -0.1934639-02 0.2184977700-02 0.4567180-01 -0.1834690-02 0.18497390-02 0.456740-01 -0.18534690-02 0.1873970-02 0.4267490-01 -0.18534690-02 0.1873970-02 0.4267490-01 -0.18734690-02 0.1873970-02 0.4267490-01 -0.18734690-02 0.1873970-02	11	-1-01099910-1-	0.40621450-01	-0.312696211-02	20-000040000000000000000000000000000000	20-019:010:0	10 10 10 10 10 10 10 10 10 10 10 10 10 1		CC328 626 D-03	-Descholation	U. 66ss7315-6.
0.2172.4762-91 -0.27111230-92 0.25471700-92 0.2779120-92 0.2779120-92 0.2779120-92 0.2779120-92 0.2779120-92 0.2779120-92 0.2779120-92 0.2779120-92 0.2779120-92 0.2779230-92 0.2779230-92 0.277770-92 0.277770-91 0.277770-91 0.277770-92 0.277770-92 0.2777770-92 0.2777770-92 0.2777770-92 0.2777770-92 0.2777770-92 0.27777770-92 0.2777770-92 0.2777770-92 0.2777770-92 0.2777770-92 0.2777770-92 0.2777770-92 0.2777770-92 0.2777770-92 0.2777770-92 0.27777770-92 0.27777770-92 0.27777770-92 0.27777770-92 0.27777770-92 0.27777770-92 0.27777770-92 0.27777770-92 0.27777770-92 0.277777770-92 0.277777770-92 0.277777770-92 0.277777770-92 0.27777770-92 0.277777777777777777777777777777777777	77	-0.12767400-14	0.36314340-01	-0.17226450-02	0. 12507 100-02	10-0-11-0-0-0-0-0-0-0-0-0-0-0-0-0-0-0-0	10-017-3083-0		0.13.60380-03		40-541055-0-04
0.27932270-01 -0.19334830-02 0.22429100-02 0.2498230-01 -0.18152250-02 0.21218230-02 0.25451730-01 -0.1254080-02 0.18471320-02 0.25478720-01 -0.1654080-02 0.184723070-02 0.2547870-01 -0.116570520-02 0.18787070-02 0.24574810-01 -0.1178500-02 0.17877970-02	13	-0.14255250-14	0.31724420-01		0.26477700-02	20.01919191919191919191919191919191919191	1010111111111		9-31533426-6.		0.44423265-03
0.21446230-01 -0.1619250-02 0.21218230-02 0.26591240-01 0.1619260-02 0.18491320-02 0.2547821-01 0.16259290-02 0.1878723370-02 0.2547820-01 0.1878790-02 0.18787990-02 0.24784810-01 -0.1859200-02 0.187879990-02 0.23478810-01 -0.185931000-02 0.18398990-02	*	-0-18851590-1-	0.27903270-01	-0.139 366 30-02	0.22429100-07	10-015 60600-01	0. 1.00 1.00 1.00		0.30/3Cacu-04	-5	0.107252376.
0.25507150-01 -0.12540660-02 0.189471320-02 0.255448720-01 -0.185480-02 0.18943070-02 0.255746800-01 -0.18593070-02 0.255746800-01 -0.185931080-02 0.18797970-02 0.25474810-01 -0.155931080-02 0.148380800-02	13	+1-0501418170-	0.27496230-01	-0.161 93250-02	0.21218230-02	-0 33934930-03	Contract of the second		*************	-0.175057500-	0-20007000000
0.25448720-01 -0.16530-02 0.18923070-02 0.4567400-01 0.74667400-01 -0.1850520-02 0.18166280-02 0.17857900-02 0.24474810-01 -0.15931080-02 0.184838080-02	16	-0-10830980-14	0.26507150-01	-0.12640660-02	0.18947328-02	-0 17730710-03	0.144.00		*********	-0.10151337-0-	0.4990004440
0.2459740U-01 -0.1870062F-02 0.1916028D-02 0.2458405U-01 -0.1716050U-02 0.191879F9F-02 0.22474810-01 -0.1593108D-02 0.14838080-02	11	-0.18385290-14	0.25476720-01	-0-10536980Y U.Z	0.18923070-02	-0 27274364-03		10-04	0.20233720-04	-0.217.05.0-05	0-13140700-05
0.24584650-01 -0.17186590-02 0.17879790-02 0.24474810-01 -0.15931080-02 0.14838080-02		-0.11324279-14	0.24597400-01	-0.18705021-02	0.1916.0200.02	20 -01 202 12.0	***************************************		40-00-tanger	-0.70s.sut13	0.35327505-05
0.23474810-01 -0.15931080-02 0.14838080-02	61	-0.14810380-14	0.24584650-01	-0.17166590-02	0.17879999-03	20.000000000000000000000000000000000000	10-024-197-7-0	10-10/10/10/-04	6.30220070-04	-0-14555-50-04	0.075-0750-05
70-000000000000000000000000000000000000	50	-0-15105650-14	0.23474810-01	-11.15031040.02	20 000000000000000000000000000000000000			*D-001001-0** - 1.0100100-04	0.49127930-04	-d. Seakastarea	0.1261030-35
				30.000.000.00	70-00000000000		0.12400046-03	40-04004001 -0.300400001010	6.14177560-04	0.14177500-0+ -0.41sec104-05	0.17950670-05

Table A-2-2 (Continued): Group 4,  $\theta$  = -0.6

Session					Moments about Mean	out Mean				
		2		,	~	•	,	•	•	10
-	-0.44630970-15	0.65526760 00	-0.3327-470 00	0.17293950 01	-0.23710930 01	0.79703670 01	-0.15793660 02	0.44204230 02	-0. 4.075010 62	0.25421171
,	-0.89595000-15	00 04126415.0	-0.36532470-01	0.15102880 00	-0.10850840 00	0.22136350 00	-0.27456/30 00	0. 477 54445 04	-0.70004470 0.3	10 01110
•	-0.11368680-14	0.128573-0 00	0.32924250-02	0.4638140-01	0.22468410-02	0.30206770-31	0-14656310-02	0.26570820-01	0.11696270-02	0 242 804 70-01
	-0.11766980-14	9.1017146, 00	0.10290590-01	0.32448240-01	0.89475520-02	0.15892300-01	0.71673950-02	0.56359860-02	0.55799490-02	0.62636143-02
•	-0-1167261 9-1+	6-74567190-51	0.47075610-02	0.18296570-01	0.40104290-02	0.81686390-02	0.34374600-02	0.50016166-02	0.24145490-02	4. 450 41 453- 67
	-0.12095680-14	0.71444220-01	0.17128033-02	0.12377620-01	0.69588330-03	0.32138610-02	0.20000000-03	0.10262490-02	0.94902180-04	0.30-60150-03
	-0.84030010-15	0.62919970-01	0.12259110-02	0.10035610-01	0.18709120-03	0.23188620-02	-0.65676760-04	0.550 30710-03	-0.73945860-0:	0. 216 67500-03
	-0.11023130-14	0.55233853-01	0.20265330-02	0.82293400-02	0.50835110-03	6.17203990-02	0.7e15Ce90-04	0.43993180-03	-0.75368720-05	0.12751583-03
•	-0.96728180-15	0.4 100e54J-01	6. 221 76380-02	0.61253030-02	0.508 151 80-03	0.10872780-02	0.13750730-03	0.22412420-03	0. 34954800-04	d. 503 784, 70-04
2	-0.90816240-15	0.42737250-01	0.12524120-02	3.48075440-02	0.33346990-03	0.82686770-03	0.83674050-04	0.17797690-03	3.22446470-94	0.43171190-04
	-0.12961850-14	0.45255000-01	0.21350640-02	0.43805620-02	0.70751240-03	0.1412/400-02	60-02646117-0	0.33649760-03	0.00-00-00-00-00-00-00-	0.11514540-01
71	-0-11882150-14	0.39316850-61	C. 337 10550-02	0.50106240-02	0.125 19780-02	0.10815550-02	0.44946010-03	G. 311 2 CRED- G3	6-14194450-61	0.101146.00-01
7	-0.86403110-15	0.35251850-01	3.16311580-02	0.36715550-02	0.62824613-03	0.62589900-03	0.21061759-03	6. 15713600-03	0. 67621930-04	0.433 90890-0
•	-0.10309910-14	0.30053450-01	0.10340670-02	0.27365760-02	0.41192020-03	0.43341830-03	0.13072750-03	0.92990629-04	0.38507040-04	0.23246220-04
2	-0.10623001.0	0.25524370-01	0.12755580-02	0.27751610-02	0.43002603-03	0.45212430-03	0.13544610-03	0.99363640-04	0.40951560-04	0.25437123-04
2	-0.12732870-14	0.2651 6430-01	0.11265430-02	0.23859070-02	0.33366540-03	0.34247550-03	0.84955700-04	0.62057820-04	0.20724630-04	0.12745630-04
11	-0.12043140-14	0-24697450-01	0.55211560-04	0.20483390-02	-0.13565200-04	0.26328630-03	-0.38990110-05	0.40045640-04	-0.80592323-05	0. 656 9 34. 40 - 05
2	-0.10075.75-14	0.23346660-01	0.16988740-03	0.19254520-02	-0.28537750-04	60-0850804-03	-0.13160130-04	0.35025460-04	-0. 55694100-05	0.54962810-04
13	-0.13397090-14	0.43412930-01	0.69054530-03	0.18121450-02	0.95502800-04	0.21514540-03	0.14735990-04	0.29860240-04	0.25026800-05	9. *** \$2169-05
62	-0.99864560-15	0.24110130-01	0.69824360-03	0.22045910-02	0.13098060-03	0.33839840-01	0.33364880-04	0.434.29090-04	0.41557150-05	0.130140-04

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Table A-2-2 (Continued): Group 5,  $\theta = 0.2$ 

Searten					Hosents about Near	OUE Meen				
		7		,	\$	9	,		•	01
	0.11124430-14	9	-0.33728970-01	0. 55455040 00	-0.12775419 00	9. 759/15eu 60	-de collection de			
~	0.11366680-14	0.21167630 00	-0.50861090-02	0.11088820 00	-0.12251521-02	3 7 14 14 14 14 14 14		200000000000000000000000000000000000000	- C. C. C. C. C. C. C.	C. 4330000000
3	0.23453460-15	0.16210930 00	0.461 90710-02	0.60470590-01	0.91384040-02		20.000000000000000000000000000000000000	0.62057670-01	0.4222047C-0	0.55227533-01
	0.22426510-15	0.10785530 00	0.45955450-02	0.31475800-01	3. 348 9067 1-02	101011111111111111111111111111111111111	10-00-10-01	0.10044210-01	0. V358 V5 LO-02	0-12653263-01
2	0.19678700-15	9.18667380-01	0.49626540-02	0.16457050-01	0.3094142000	10-010-010-01	0	0.00502310-02	つうろうりゅうへいてつ	0.400 reserved
	6.28282930-19	0.61566170-01		10167670	0 46716770 01	20-01770-10-00	0.16.11.10-02	0-14445450-02	0.0.000000000	0.000110000
1	0.267-2530-15	0.52458143-01	0. 88812710-03	4 40.185190-07	0 269 374 7 1 03	30-011 5137300	0-17/17/2010-03	0.57065055-03	0.005-303-04	0.10555510-03
	6.271172710-15	0.47775563-01	6. 147.28 x10-02	0 42640340-02	50-01900000	0.23101010	0.326/2020-00	0.20530000-03	0.34659710-04	0-10455400-04
	0.25965340-15	3.41346210-01	6.25383660-02	0 48101720-02	0 117 423 73	0.10007-10-02	0.0144416.0	0.21750720-03	C. I. Seules J-C.	0.54447440-04
	0.24605320-15	0. 35044.70-01	6.44001630-02	20 4311741410	70-0177711110	0.100,000,000	0.2416/625-05	0.20+72910-03	0.15276282-03	0.5075955000
***	0.23051910-15	0.34855073-01	0.29964740-02	0 33561500-03	20 -00 701 111 10	70-1700000110	0-100220001-03	0.48867680-03	0.25205210-0,	0-15700643-03
7	9.21566080-15	0.30303740-01	0.227 81613- 02	0 22950190-02		200444717-03	9	0.12622140-03	0.55% 1007-64	0.32*******
•	0.21593840-15	0.26095490-01	Q. 1640×580-02	0.16954220-02	10-101101101	6.4.2007.07.03	0.000/00000	0.34431855-04	0.12395240-04	6.5vvelili-05
	9.21219140-15	0.26509230-01	6. 10461190-02	6.16709700-02	20-173-173-0	00.101/10430-03	***********	0.21998350-04	0.14702653-02	0.34072750-05
- 5	3.32612830-15	0.23680950-01	0.77414210-03	0.12524116-02	0.32004813-03	0 1000000000000000000000000000000000000	0. 2 627 1 1 1 04	0.14703180-04	0-5705-562-0	0.16762143-05
	0.38344330-15	0.21247660-01	0.47145850-03	0.19493:9(4-02	20-0-0-0-0-0-0	100000000000000000000000000000000000000	0.134.0/20-04	0.63305500-05	0-153077av-05	0.66000980-06
1	0.27686190-15	0-19606350-01	0.1855779.403	0.86611170-03	200000000000000000000000000000000000000	0.10021020	0.003001co-05	0.57255000000	0.9194200,-00	0.5450445400
	0.40328850-15	0.18791660-01	0. 705 3712 04	0 4798 4990-03	20 -01 707 27 10	10-012/12/00	00-001007000	0.40565640-05	0.04462773-01	0.3413622-06
6	0.33181790-15	0.18296270-01	0.13585893-03	0.771.11550-01	00-000000000000000000000000000000000000	0.300.300.00	-3.39366.20-07	0.44020730-05	-0.1.77651,-07	0. 467 286 30-06
	0.49601320-15	0.16801750-01	0.14683840.03	0 71061500-03		50-1001-00	0.10410650-05	0.40768530-05	0.25435550	0.37944745-04
				0.11031310-03	0-13104110-04	*********	0.10010720-06	0.35177970-05	G. 13503610-07	6-3115-0 pu-0e

Table A-2-2 (Continued): Group 6,  $\theta = 1.0$ 

200					Moments about Mean	out Mean				
	1	2	3	,	•	•	1	*	•	2
,	D.L. Backback	3.71.017.0 00	0.402.8110 00	O. telusciote	0.21064070 01	0.64642620 01	0.93907550 01	0.25132290 02	0.35890720 62	u.15328223 03
7	3	00	0.84214150-11	0.363c3c3c	0.36323670 00	0.90837950 00	0.15537900 01	0.34165650 01	0.66305450 01	0.1194137. 32
•	0.16.000000-14	00 02621661.0	G.2861263J-02	0.23050530-01	0.55+71550-02	0.42225790-01	0.51381580-02	9.37344380-01	0. 45157260-02	U. 367+6323-91
	J. 152130su-14	J. 00. 13 [ 30-01	-0-2-030000-02	0.2.900.001	-0.16774403-02	0.10296:20-01	-0-91814090-03	0.56267260-02	-0.67273910-03	0.346101337-02
•	0.1707+590-14	3.13.203101	-0.21635326-02	0.10300661-01	-0.1255e10J-02	0.52755590-02	-0.72335540-03	0.20572950-02	-0.39985830-03	0. 336 -4743-01
	J.16.00.050-1.	Je25.14.4.0-01	-0-84 //3350-04	0.41226240-04	0.3483170,-04	0.17952470-32	0.65145950-04	0.48427220-03	0.41333700-04	0.145 30.66.3-31
	3-135-50-14	1.40130012-01	-0.37941355-02	0.00301230-02	-0.13614620-02	0.11576610-02	-0.40517940-03	0.27039650-03	-0.11651350-03	0-164411 30-0
	3.106/5233-1-	0	-0.25624010-02	6.31612040-04	-0.51504220-03	0.85467720-03	-0.79682560-04	0.16966460-03	-9. 69 280260-05	0.17235153-04
	3.137cc776-1.	Deserve Iso-01	-0.2-51 7845-04	0.35701.30-02	-0.53771433-03	0.56036630-03	-0.10794350-03	0.92787700-04	-0.22075470-04	0.16954933-04
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7.5	4.165.01.6.14	0.31/111430-01	-0.12358130-03	0.336/9330-02	0.43997410-04	0.56109010-33	0.51238320-04	0.11505170-03	0.20241200-04	0.26598943-94
17	0.12+50/35-14	10-01/12/12/20	-0.92517e7u-04	0.2686+000-02	-0.65340510-05	0.39384260-03	-0.12468770-05	0.68255720-04	-0.61804829-07	0.12663690-04
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12	0.14-07505-1-	0-10-10010-01	-0.44185310-03	50-00004162.5	-0.20301720-03	0.48502350-03	-0-12358100-03	0.11909070-03	-0.45126610-04	0.33593120-04
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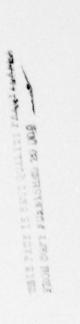


Table A-2-2 (Continued): Group 7, 0 = 1.8

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Table A-2-2 (Continued): Group 8, 8 = 2.6

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